

Leveraging statistical learning theory to characterize the U.S. water consumption

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Key Points:

- Statistical inference of the nation wide water withdrawal patterns in the U.S.
- Irrigated farming, thermoelectric energy generation and urbanization are the most water-intensive anthropogenic activities
- Water withdrawal patterns across U.S. show varying sensitivity (between $\pm 10\%$) to future changes in precipitation changes under the RCP8.5 scenario.

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Abstract

Access to accurate estimates of water withdrawal is requisite for urban planners as well as operators of critical infrastructure systems to make optimal operational decisions and investment plans to ensure reliable and affordable provisioning of water. Furthermore, identifying the key predictors of water withdrawal is important to regulators for promoting sustainable development policies to reduce water use. In this paper, we developed a rigorously evaluated predictive model, using statistical learning theory, to estimate state-level, per-capita water withdrawal as a function of various geographic, climatic and socio-economic variables. We then harnessed the data-driven predictive model to identify the key factors associated with high water-usage intensity among different sectors in the U.S. We analyzed the predictive accuracy of a range of parametric models (e.g., generalized linear models) and non-parametric, flexible learning algorithms (e.g., generalized additive models, multivariate adaptive regression splines and random forest). Our results identified irrigated farming, thermo-electric energy generation and urbanization as the most water-intensive anthropogenic activities, on a per-capita basis. Among the climate factors, precipitation was also found to be a key predictor of per-capita water withdrawal, with drier conditions associated with higher water withdrawals. Results of the first-order sensitivity analysis indicated changes between $\pm 10\%$ in the future water withdrawal across the U.S., in response to precipitation changes, by the end of the 21st Century under the business-as-usual scenario. Overall, our study highlights the utility of leveraging statistical learning theory in developing data-driven models that can yield valuable insights related to the water withdrawal patterns across expansive geographical areas.

1 Introduction

Integrated water resource management has been receiving increasing attention globally (Giordano & Shah, 2014; Rahaman & Varis, 2005). Rapid growth in population, and increased rates of economic development and urbanization have resulted in increased demands for fresh water in energy, agriculture, industry, and the commercial and residential sectors, all of which have severely stressed water resources in many regions. Sustainable management of demand for water has been brought into the limelight in the United States following several devastating, multi-year drought episodes in California and the Midwest which led to adverse impacts on agricultural productivity and energy generation capacity, costing the U.S. economy tens of billions of dollars. According to the U.S. Environmental Protection Agency, 40 out of 50 states will expect water shortages in some portion of their jurisdiction in the next 10 years, even under average conditions (EPA, 2017).

Accurate estimates of short-, medium-, and long-term demand for water is valuable for urban planners, regulators and operators of critical infrastructure systems to ensure reliable and affordable provisioning of many critical services including water. Optimal investments in the design, operation, modernization and expansion of water infrastructure systems are largely dependent on access to realistic and credible predictions and projections of the spatio-temporal variability in demand for water (Billings & Jones, 2008). According to Hall, Postle, and Hooper (1989), “the success of any water resource development is critically dependent upon the reliability of the forecasts of future water demands that are employed in its design (and management)”.

In this paper, we leverage statistical learning theory to: a) develop accurate predictive models for per-capita water use in various sectors in the U.S., b) identify the key predictors of state-level, per-capita water withdrawal, c) understand the relationship between each of the key predictors and per-capita water use, and d) analyze the sensitivity of the water withdrawal patterns to changes in climate variability (e.g., precipitation changes) under changing climate conditions. Our predictive water withdrawal models were developed using state-level, per-capita water withdrawal data over the past two decades

65 – together with various geographic, climatic, and socio-economic factors – to identify the
66 key factors that are associated with high water-usage intensity among different sectors
67 in the U.S.

68 We hypothesized that statistical models that assume ‘rigid’ functional forms – such
69 as linearity and additivity (e.g., multiple linear regression) – would not adequately cap-
70 ture the complex dependencies between state-level water withdrawals and socio-economic
71 and geoclimatic conditions; and that more robust statistical learning algorithms (e.g.,
72 ensemble-of-trees), would be more effective in predicting state-level, water withdrawals.
73 Moreover, given that the largest fraction of water-withdrawals occur in the agricultural
74 and thermoelectric generation sectors, we hypothesize irrigated farming and power gen-
75 eration to be the key predictors of state-level water withdrawals.

76 The structure of this paper is as follows. The review of the existing literature in
77 predicting water withdrawal is summarized in Section 2. Data and methods are intro-
78 duced in sections 3 and 4, respectively. Results are summarized in Section 5, followed
79 by the concluding remarks in Section 6.

80 2 Background

81 A plethora of research studies have focused on analyzing, predicting and project-
82 ing water demand – with various different spatio-temporal scales and lead time-horizons
83 – using a range of methods such as simulation, econometrics and statistical learning the-
84 ory. Donkor, Mazzuchi, Soyer, and Roberson (2014) reviewed research articles on wa-
85 ter demand forecasting – published between 2000 and 2010 – to identify useful models
86 for water utility decision making. They concluded that artificial neural networks were
87 more popular for short-term demand-forecasts, while econometrics, scenario-based and
88 simulation models were more likely to be used for making long-term strategic decisions.
89 They also highlighted the value in probabilistic forecasting to capture uncertainties as-
90 sociated with future demand. More recently, Seabri (2016) surveyed the empirical liter-
91 ature on urban water forecasting using a meta-analytical approach. Their meta-regression
92 analysis concluded that model accuracy depended on the scale of analysis, the type of
93 approach used, model assumptions and sample size. Hamoda (1983) examined the im-
94 pact of socio-economic factors on the residential water consumption in Kuwait. More specif-
95 ically, Hamoda (1983) leveraged linear regression to characterize the impacts of income,
96 market value of land, rents of dwellings and household size on average per-capita water
97 consumption. They concluded that the hot climate of Kuwait together with its contin-
98 ually improving standards of living were the primary factors contributing to high wa-
99 ter consumption rates in the country.

100 In an another study by Lutz et al. (1996) leveraged a variation of the EPRI (Elec-
101 tric Power Research Institute) model to study the patterns of residential hot water con-
102 sumption. Their study shed light on the impacts of efficiency standards for water heaters
103 and other market transformation policies. Jorgensen, Graymore, and O’Toole (2009) an-
104 alyzed the social factors in residential water-use and highlighted the importance of inter-
105 personal and institutional trust for implementation of effective water conservation schemes.
106 Sovacool and Sovacool (2009) implemented a county-level analysis of the energy-water
107 nexus in the U.S., and concluded that twenty-two counties will likely face severe water
108 shortages, brought about primarily due to increased capacity expansion in thermoelec-
109 tric generation. Chandel, Pratson, and Jackson (2011) leveraged a modified version of
110 the U.S. National Energy Modeling Systems (NEMS) together with thermoelectric water-
111 use factors from the EIA to investigate the impact of various climate change policy on
112 the energy mix. They found that all of the climate policy scenarios that were considered
113 in the study could lead to a reduction in fresh water withdrawal for power generation,
114 compared to the business as usual scenarios. Moreover, they found that water-use de-
115 creased as the policy’s carbon price increased. Davies, Kyle, and Edmonds (2013) lever-

116 aged GCAM – an integrated assessment modeling of energy, agriculture, and climate change
 117 – to assess the water intensity associated with electricity generation until 2095. They
 118 found that water use would likely decrease with capital stock turnover.

119 The majority of the empirical studies to date have focused primarily on either a
 120 particular geographical location, or a given sector in the U.S., and leveraged either lin-
 121 ear models (the assumptions of which may not be supported by the empirical data) or
 122 ‘black-boxes’ (e.g., artificial neural network) to project demand. This paper will use state-
 123 of-the-art statistical learning techniques to analyze water withdrawal data – available
 124 from USGS over the past two decades for the entire U.S. – and develop an accurate and
 125 interpretable predictive water withdrawal model as a function of socio-economic, geo-
 126 graphic, climatic conditions.

127 It is noteworthy that, though not pursued in this study, there exist another fun-
 128 damentally different approach to modeling water withdrawal, based on complex, mech-
 129 anistic hydrologic models with integrated elements of human-water interfaces (e.g., Pokhrel,
 130 Hanasaki, Wada, & Kim, 2016; Wada et al., 2017). Models in this category include, for
 131 instance, PCR-GLOBWB (Sutanudjaja et al., 2018; Wada, Wisser, & Bierkens, 2014),
 132 WaterGAP (Alcamo et al., 2003; Flörke et al., 2013), and H08 (Hanasaki et al., 2008a,
 133 2008b). These models have varying ranges of processes accounting for the coupled hu-
 134 man and natural systems. Despite the utility of these models in providing a mechanis-
 135 tic understanding on the functioning of the system, they are inherently complex and dif-
 136 ficult to parameterize – partly owing to the limited availability of observational data-
 137 sets. Different sorts of simplifications and conceptualizations are therefore necessary to
 138 model the complex interactions between human and natural systems (e.g., Wada et al.,
 139 2017). Our proposed modeling paradigm – based on statistical learning theory – can be
 140 complementary to hydrological modeling efforts. Our approach offers key advantages of
 141 a) being computationally efficient, and b) requiring a limited set of predictors to re-construct
 142 the continuous space-time evolution of water withdrawal; which can be used to fur-
 143 ther constrain the parameterization of more complex, mechanistic hydrologic models. In
 144 summary, our approach can help identify the most water-intensive sectors across vari-
 145 ous states, inform policy makers, regulators and researchers on the exiting U.S. water
 146 use patterns and identify sectors and areas where efficiency and conservation mechanisms
 147 could yield maximum return, in-terms of enhanced sustainability of our urban ecology.

148 **3 Data and Initial Analysis**

149 Data were collected from various publicly available sources such as the Geological
 150 Survey website (USGS, 2017), the Energy Information Administration (EIA, 2017), the
 151 Bureau of Economic Analysis (BEA, 2017), the U.S. Census Bureau (USCB, 2017), the
 152 Climate Prediction Center (CPC), the National Weather Service (NOAA, 2017), the U.S.
 153 Department of Agriculture (USDA, 2007), the Coastal States Organization (CSO, 2017),
 154 the U.S. Environmental Protection Agency (EPA, 2017) and other sources (IOWA, 2017).
 155 Below, we will provide a brief description of our response variable (i.e., per-capita wa-
 156 ter usage) and various socio-economic, hydro-climatic and geographic predictors that were
 157 used in our analyses. It should be pointed out that since the water withdrawal data is
 158 only available at five-year increments, the predictors were processed to match the tem-
 159 poral scale of our response variable.

160 **3.1 Response Variable: Per-Capita, State-Level Water Withdrawal**

161 State-level water withdrawal data (in million gallons per day) were selected as our
 162 response variable, and were obtained from U.S. Geological Survey website (USGS) for
 163 the period of 1991-2010. USGS water usage data are collected and compiled every five
 164 years for each of the 50 states, the District of Columbia, Puerto Rico, and the U.S. Vir-
 165 gin Islands. The data source provides a breakdown of water usage in eight different sec-

166 tors (depicted in Fig. 1) such as thermoelectric, irrigation, public supply, industry, aqua-
 167 culture, domestic, livestock and mining. Thermoelectric and irrigation are the two dom-
 168 inant sectors that account for almost two-third of the total water withdrawal across the
 169 U.S. We, however, note that there is a large regional variability in water withdrawal pat-
 170 terns – the States in the east is more dominated by the thermoelectric and industrial wa-
 171 ter sectors, while the irrigation is the main water usages in the central and western part
 172 of the U.S. To control for the varying sizes of states, we normalized the state-wide to-
 173 tal water withdrawal data by the total population of each state. The distribution of state-
 174 wise, normalized water withdrawal for years of 2006–2010 can be seen in Fig. 1(bottom
 175 panel). States highlighted in shades of red represents high per-capita water usage, while
 176 the states in blue represent low per-capita water usage. Fig. 1(bottom panel) reveals that
 177 Idaho has the highest per-capita water usage for the year 2006–2010.

178 The distribution of the per-capita water withdrawal (in million gallons per day)
 179 for the period 1991-2010 is depicted in Fig. 2. The distribution of per-capita water with-
 180 drawal is right-skewed and has a heavy-tail distribution. In fact, it can be seen that the
 181 power-law distribution provides a reasonable fit to the tail of the data (red line in Fig. 2a).
 182 Power-law distributions describe phenomena where large events are quite rare, but small
 183 events are very frequent. Fig. 2 suggests that a small fraction of the states in the U.S.
 184 tend to consume disproportionately large volumes of water per capita.

185 **3.2 Socio-Economic Predictors**

186 Gross State Product (GSP) data were collected from the U.S. Bureau of Economic
 187 Analysis for the years of 1991-2010 in current value. The GSP data (in millions of USD)
 188 were then converted to time value of 2010, using the GDP deflator. Household Median
 189 Income (in USD) was collected from the Bureau of Labor Statistics. The value of income
 190 data was converted to 2013 CPI-U-RS (Consumer Price Index Research Series Using Cur-
 191 rent Methods) USD.

192 The education level data obtained from the U.S. Census Bureau contains the fol-
 193 lowing four levels for each reported year: (a) percentage of population with less than high
 194 school diploma, (b) percentage of population with high school diploma only, (c) percent-
 195 age of population some college (1-3 years), and (d) percentage of population with four
 196 years of college or higher. We leveraged generalized additive models to impute the miss-
 197 ing data and align the temporal scale of the education data with that of water withdrawal.
 198 The premise for including this variable in the analysis is to test whether educational lev-
 199 els are predictive of the public supply water withdrawal.

200 Datasets related to thermoelectric energy generation – e.g., coal, petroleum, and
 201 gas fired plants, nuclear and geothermal technologies – in mega watt-hours were collected
 202 from the Energy Information Administration (EIA). Coal production, available from the
 203 EIA, was used as a proxy for mining industry, since coal is the biggest profit generat-
 204 ing mining production in the U.S. The percentage of urban population data were col-
 205 lected from the U.S. Census. Since the temporal scale of the urban population data were
 206 decadal, the years did not match the years in the USGS water dataset. We therefore im-
 207 puted the missing years of the percentage of urban population data a using generalized
 208 additive model to match the years across the two datasets.

209 **3.3 Hydro-climatic and Geographic Predictors**

210 Time-series of datasets related to Cooling Degree Days (CDD) and Heating De-
 211 gree Days (HDD) are based on variation in air temperature estimates which were made
 212 available from Climate Prediction Center (CPC) and National Weather Service (NWS).
 213 Other hydro-climatic variables as predictor variables include Standardized Precipitation
 214 Index (SPI), soil moisture, and annual precipitation data were provided by the National

Centers for Environmental Information. The SPI characterizes the inter/intra-annual variability of precipitation with positive values indicating wetter than normal conditions and the negative values being indicative of drier than normal conditions (Hayes, Svoboda, Wall, & Widhalm, 2010; McKee, Doesken, & Kleist, 1993). Additionally, we used the upper 1 m simulated soil-water content (mm) based on the CPC model based simulations to represent the near-surface wet and dry conditions (see Fan & van den Dool, 2004, for more details).

Coastal status was calculated for each state by creating dummy variables indicating whether the state is in the borders of (a) the Atlantic Ocean, (b) the Pacific Ocean, (c) the Gulf of Mexico, and (d) the Great Lakes. The states in proximity of any of the above-mentioned water-sheds, were coded as '1', and otherwise as '0'. The estimates of the total irrigated farmland area were collected from the Census of Agriculture Farm and Ranch Irrigation Survey (2008), conducted by the National Agricultural Statistics Service (NASS) in the U.S. Department of Agriculture (USDA). The surveys are conducted every five years, starting from year 1992. To align the time steps of the farm data with that of water usage, we used data from 1992 to represent irrigated farmland size between 1991 and 1995, and 1997 data was used to represent the value between 1996–2000. We normalized the data by the total land size of each state to obtain the percentage of irrigated farmland area per state. Prior to the analysis and the model set-up, all predictor variables were aggregated spatially and temporally to match the state-wide, five-yearly available water withdrawal datasets.

3.4 Exploratory Data Visualization and Analysis

A 'biplot' is a useful visualization tool for multivariate data. One of the most commonly used types of a biplot is based on principle component analysis. A PCA-biplot is a low-dimensional representation of multivariate data, using only the first two principle components. In a PCA-biplot, vector lengths approximate standard deviations, and the cosines of their angles are proportional to the correlation between the variables. It can be seen from Fig. 3 that over the years of 1995–2010, the state-level water usage did not change significantly. For example, on the bottom left corner of the plot, we observe that water usage of Arizona, Louisiana, Texas, and Florida are located close to each other across the different years. The energy generation and cooling-degree-days (CDD) vectors extended in the direction of Texas suggest that the state's thermoelectric power generation and its hot climate can help explain the variance of water usage in Texas, as opposed to states of Colorado or North Dakota which lie close to the heating-degree-day (HDD) vector. Moreover, the Fig. 3 reveals that while water usage in the densely populated states of the Northeast can be explained by socio-economic factors such as income and education and measures of urbanization, the water usage in the larger Midwestern and Western states of North and South Dakota, Nebraska, Iowa and New Mexico tend to be dominated by farming and mining practices.

4 Methodology

The existing empirical literature in field of water analysis has almost exclusively focused on descriptive and explanatory statistical modeling, while predictive modeling of water analysis has largely been under-explored. Unlike descriptive or explanatory modeling which is concerned with best explaining the past variability in the data, predictive modeling is concerned with predicting 'new/unseen' data. The expected prediction error (*EPE*) for a new observation x can be summarized by the equation below [11]:

$$\begin{aligned}
EPE &= E \left[Y - \hat{f}(x) \right]^2 \\
&= E \left[Y - f(x) \right]^2 + \left[E \left(\hat{f}(x) \right) - f(x) \right]^2 + E \left[\hat{f}(x) - E \left(\hat{f}(x) \right) \right]^2 \\
&= Var(Y) + Bias^2 + Var \left(\hat{f}(x) \right)
\end{aligned} \tag{1}$$

261 The first term represents the irreducible error which is the result of the inherent
262 stochasticity in any process. The second term (the bias) represents how closely the es-
263 timated function mimics the process of interest, and the third term (variance) arises due
264 to using (noisy) samples to estimate the response function. Descriptive and explanatory
265 statistical models often focus on reducing the bias of the estimate. However, predictive
266 modeling focuses on minimizing the bias and variance *simultaneously*. The central the-
267 sis in this paper is that, with the recent accelerated pace of large complex datasets be-
268 coming available, predictive modeling can be leveraged as a powerful tool to identify com-
269 plex and non-linear dependencies that can lead to generating new hypothesis and ad-
270 vance the scientific discovery in the field.

271 In the next section, we will present a brief discussion on supervised learning the-
272 ory and predictive modeling. We will then present a detailed discussion of the algorithm
273 that was used to develop the final best predictive model of the state-level, water with-
274 drawal data.

275 4.1 Supervised Learning Theory (Predictive Modeling)

276 Supervised learning theory was leveraged to develop accurate predictive models for
277 state-level water withdrawals, and identify their most important predictors of in the U.S.
278 The main objective of supervised learning is to approximate a process of interest (e.g.,
279 water withdrawals) as a function of various independent predictors (e.g., geographic, cli-
280 matic and socio-economic factors). Mathematically, the prediction process can be sum-
281 marized by $y = f(X) + \epsilon$; where the stochastic additive Gaussian noise ϵ represents
282 the dependence of y on factors other than X that are not controllable. The goal of su-
283 pervised learning is to leverage the observed records and approximate the response $hat{f}(X)$
284 (i.e., water withdrawal) such that the loss function L is minimized over the entire do-
285 main of the input data space:

$$286 \quad L = \int w(X) \Delta \left(\hat{f}(x), f(x) \right) dX \tag{2}$$

287 where $w(X)$ is a possible weight function, and Δ represents the Euclidean distance
288 (or other measures of distance). The value of L in the equation above characterizes the
289 accuracy of the estimate over the entire domain (Hastie, Tibshirani, & Friedman, 2009).

290 We trained our data with various parametric (e.g., generalized linear models) and
291 non-parametric (e.g., generalized additive models (GAM), multivariate adaptive regres-
292 sion splines (MARS) and random forests (RF)) methods – description of which can be
293 found in the Appendix. Given that the ensemble tree-based algorithm (the method of
294 random forest) outperformed all other algorithms in terms of out-of-sample predictive
295 accuracy (see Section 5), we selected it as our final best model. A brief description of
296 the random forest (RF) algorithm is provided below.

297 4.2 Random Forests (RF)

298 Random Forest is an ensemble decision tree-based method developed by Breiman
299 (2001), and can be mathematically represented as:

$$F(x) = \frac{1}{m_{\text{tree}}} \sum_{i=1}^{m_{\text{tree}}} T_i(x) \quad (3)$$

where T_i is a single decision tree, trained on bootstrap samples from the original data and x represent a p -dimensional vector of input data predictors (e.g., the geographic, climatic and socio-economic factors used in this analysis). The subset of predictors for building each decision tree is randomly selected, and best splits values are chosen such that the sum of squared errors (or least absolute deviation) within each node t within T_i is minimized. Each decision tree is developed by recursively splitting the data space into terminal nodes, until each terminal node contains no more than a certain predefined minimum number of records. The average (or mode value as for the case of classification) is then assigned to the terminal nodes. $F(x)$ estimates the response value, by aggregating m such decision trees.

Regression trees are low in bias, particularly if they are grown sufficiently deep, since the tree structure follows the structure of the data well so that the estimated target mean is close to the true mean (Hastie et al., 2009). They are, however, notoriously noisy, and generally have high variance. They are unstable and not particularly robust to outliers, and this makes the procedure non-ideal for datasets that contain many outliers. The issue of high variance is solved by leveraging the ensemble methodology as a variance reduction technique. The ensemble-of-trees methods such as random forest are generally very robust to outliers and offer strong predictive power. The estimation of prediction error of random forest can be obtained by leveraging the out-of-bag (OOB) data (i.e., the test data that was set aside during the development of each tree and not used in building that tree) to compute the mean square error as below:

$$MSE = \frac{1}{n} \sum_{i=1}^n (y_i - \bar{y}'_i)^2 \quad (4)$$

where \bar{y}'_i is the average OOB predictions data for the i^{th} observation (Liaw & Wiener, 2002). Since the method of random forest is non-parametric, partial dependence plots (PDPs) can be used to implement variable inference. PDPs calculate the marginal effects of a given predictor variables x_j in a “ceteris paribus” condition (i.e., controlling for all the other predictors). Mathematically, the estimated PDP is given as (Hastie et al., 2009):

$$(\hat{f}_J)(x_j) = \frac{1}{n} \sum_{i=1}^n (\hat{f}_J)(x_j, x_{-j,i}) \quad (5)$$

where \hat{f}_J is the approximation of the true function that generates y ; n is the size of the response vector (i.e., the size of the training dataset); x_{-j} represents all input variables except x_j . The estimated PDP of the predictor x_{-j} provides the average value of the function \hat{f} when x_j is fixed and x_{-j} varies over its marginal distribution.

5 Results and Discussion

Table 1 summarizes the performance of each of the models. The first column summarizes the goodness-of-fit for each of the models. Multivariate adaptive regression splines (MARS) and the method of random forest (RF) fit the data substantially better compared to multiple linear regression (MLR) and generalized additive (GAM) model. The second and third columns in Table 1 show the in-sample and out-of-sample root mean squared errors for each of the models. Again, it can be observed that MARS and RF are

341 competitive in terms of in-sample fit, but RF significantly outperforms all other mod-
 342 els, in terms of out-of-sample accuracy. In fact, the analysis of variance test on the pre-
 343 diction errors of the different models revealed statistically significance differences between
 344 the mean errors, with a p-value $< 2 \times 10^{16}$.

345 Fig. 4 (top panel) visualizes the fit of each of the prediction models. The predic-
 346 tion model based on the random forest algorithm substantially outperforms all other mod-
 347 els in terms of the goodness-of fit. The model developed using the random forest algo-
 348 rithm was therefore selected as the final best model.

349 In order to further demonstrate the predictive capability of the model, we trained
 350 the random forest algorithm with the data until the end of 2005 in order to predict wa-
 351 ter withdrawals in an independent testing period of 2006–2010. Table 2 summarizes model
 352 fit and predictive accuracy, and Fig. 4 (bottom panel) provides a graphical representa-
 353 tion of the predicted and observed values of per-capita water withdrawals. Based on the
 354 results summarized in the table and the plot, it can be inferred that RF outperforms all
 355 other models. In fact, RF is able to estimate the water usage above 5 million gal/day/person
 356 accurately, even though there are less observation points. While MARS performs well
 357 below 5 million gal/day/person (where there is more observations) it performs poorly
 358 where the data is sparse.

359 These results confirms our hypothesis that simple linear-based models (e.g., MLR)
 360 and additive structures such as GAMs are not able to capture the complex relationships
 361 in the data adequately. Moreover, the fact that RF outperformed MARS is not surpris-
 362 ing. MARS can be seen as an extension of recursive partitioning algorithms such as tree-
 363 based methods (Friedman, 1991) which is very effective at capturing high order inter-
 364 actions and yielding low-bias estimates. However, the model is not as effective in vari-
 365 ance reduction and therefore has an inferior predictive power.

366 We leveraged a data-driven variable selection, based on an algorithm proposed by
 367 Genuer, Poggi, and Tuleau-Malot (2010), to implement input variable reduction for the
 368 RF model. The variable selection algorithm first involved developing multiple forests and
 369 ranking their input variables (based on their importance by calculating their contribu-
 370 tion to out-of-sample predictive accuracy, and their standard deviations). Variables at
 371 the bottom of the list (in terms of importance) whose standard deviation was below the
 372 minimum calculated threshold were removed. Multiple nested models were then devel-
 373 oped in a step-wise forward strategy. The smallest subset of input data that yielded the
 374 best predictive accuracy were retained for the final model. The list of the final key vari-
 375 ables selected for each sector are shown in Fig. 5.

376 The importance plot shows the ranking of the variables in terms of their contri-
 377 bution to the model’s out-of-sample predictive performance, with the variable highest
 378 on the y-axis contributing the most to model’s performance. It can be observed that the
 379 percentage of irrigated farmland is the most important predictor of state-level per-capita
 380 water withdrawal, followed by total state-level precipitation, heating degree days (HDD),
 381 urbanization, thermoelectric energy generation and state-area. This result is intuitive,
 382 since irrigation and mining generally comprise a large share of water withdrawal in the
 383 U.S.

384 In order to understand the association between the top most important predictors
 385 and our response variable (per-capita water withdrawal), partial dependence plots were
 386 examined. Below, we will discuss the partial dependencies for each of the predictors, in
 387 order of their importance ranking depicted in Fig. 5.

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5.1 Effect of Percentage of Irrigated Farmland Areas

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The partial dependence between the percentage of irrigated farmland and per-capita water withdrawal indicates a positive association, with larger irrigated farmlands being associated with higher water withdrawal intensity. This is intuitive, as the U.S. agricultural sector accounts for a significant fraction of total water consumption. Some of the states associated with the different percentiles of water withdrawal have been highlighted in Fig. 6. As expected, states such as Nebraska and Arkansas lie at the extreme right end of the graph due to their large irrigated agricultural lands. Nebraska is ranked first in the U.S. in terms of total irrigated acres of land, and has seen rapid expansions of irrigated farmlands in recent years. It is located on the Ogallala Aquifer which is among the largest in the world, and makes heavy use of ground water for farming and irrigation. In fact, most of the irrigation in Nebraska (and effectively all of the more recent expansion in irrigated farming) is pumped from the High Plains (aka Ogallala) Aquifer. Arkansas, the number one producer of rice in the U.S., also lies at the extreme right end of the table, which is not surprising since rice is among the most water-intensive crops (Johnson, Christopher, Anil, & NewKirk, 2011). It is interesting to note the step-function jump from the states such as Delaware to the state of California. This could suggest that the crops grown in Delaware that are mostly corn, soybeans and wheat-based may be less water intensive than the crops grown in CA (mainly nuts, and fruits).

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5.2 Effect of Precipitation Variability

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We hypothesized higher precipitation levels to be associated with decreased water usage since precipitation affects a variety of sectors such as thermoelectric power generation, irrigation, public supply, industry, aquaculture, domestic, and life stock. The observed pattern in Fig. 6 is consistent with our initial hypothesis, indicating that wetter regions use less water. However, the decreased water-use plateaus at the threshold of 700 mm of precipitation

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5.3 Effect of Heating Degree Days

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Heating degree days (HDD) measure the difference between average air temperature and an arbitrarily chosen standard baseline temperature (typically 65°F in the US) to which the built environment would be heated on cold days. Annual HDD measures the time-integrated variation over a year between the average daily temperature and the baseline 'comfort' temperature. Interestingly, there seems to be a subtle, positive association between heating degree days and water withdrawal, with a sudden jump past HDD of 3000 which is mostly associated with the states located in the North-Central parts of the U.S., such as North Dakota, Minnesota, Wyoming and Montana (Fig. 6). This might be attributable to the (non-coal) mining and industrial activities such as fracking in these northern states. For instance, in 2005, Minnesota had the largest share of (sulfide) mining-related fresh water withdrawals in the U.S. Wyoming and Montana also have an active mining sector. Moreover, a significant amount of water is used in North Dakota in hydraulic fracturing for oil and gas. Unfortunately, data limitation as well as the diversity and rapid shifts in these mining and fracking activities make it difficult to test these hypotheses.

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5.4 Effect of Percentage of the Urbanized Areas

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The partial dependency plot for the urbanization effects on water withdrawal patterns across U.S. clearly shows that the more urbanized states tend to be less water-intensive (Fig. 6). Again, this is largely due to the fact that the domestic sector and public supply sector comprises a significantly smaller fraction of total water withdrawal as compared to the farmland or energy generation sectors.

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5.5 Sensitivity of Water Withdrawal to Future Climate Variability

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In this section, We demonstrate the utility of leveraging the predictive model, based on the random forest algorithm, in assessing the sensitivity of changes in water withdrawal patterns across U.S. in response to changing climate conditions. To this end, we used the precipitation datasets from the five CMIP5 Global Circulation Models (GCMs: HadGEM2-ES, IPSL-CM5A-LR, MIROC-ESM-CHEM, GFDL-ESM2 and NorESM1-M), available in a bias-corrected form by the Inter-Sectoral Impact Model Intercomparison Project (ISI-MIP; Warszawski et al., 2014, see also www.isimip.org for more details). For this demonstration purpose, we aggregated the daily precipitation dataset to create state-wide, mean annual estimates for the two time periods indicating the contemporary condition (1995-2010) and the future one (2070-2085), which are taken from the runs corresponding to the RCP8.5 future pathways under the narration of a “business-as-usual” scenario. For these periods, we run the established RF model to predict state-wide water withdrawal using their respective precipitation data-sets while keeping other variables at nominal values following a “ceteris paribus” condition. We estimate the ensemble mean of the state-wise, projected changes in the water withdrawal rates based on the RF model outputs driven by five GCM based precipitation data-sets.

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We observed a clear north-south gradient in the relative changes of the water withdrawal patterns across U.S. between future and contemporary period estimates (Fig. 7). Our simulation results indicated increased water withdrawal rates in the southern States, while the declined rates are expected in the Northern states – in response to future precipitation changes. The southern states such as Texas (TX), Florida (FL), Louisiana (LA), and Arizona (AZ) show a projected increase of more than 5% in their water withdrawal rates relative to the contemporary condition. The changes in the future water withdrawal rates across the majority of States is in-between $\pm 10\%$ with the driving precipitation changes being projected $\pm 15\%$. Results of this analysis also indicate a varying level of sensitivity in the projected water withdrawal rates to changes in precipitation estimates (Fig. 7; bottom scatter plot). For example, in states such as Texas (TX) and Arizona (AZ), a small change in mean annual precipitation (around 2%) creates a relatively larger change in water withdrawal (6-8%). Notably, all of the above presented estimates corresponds to ensemble mean of the modeled water withdrawal (based on the RF model run with five GCMs outputs); analysis based on the individual model estimates revealed a substantial uncertainty owing to the differences in projected precipitation from different GCMs.

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6 Conclusions

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In this paper, we analyzed the predictive accuracy of various statistical methods in predicting the state-level, per-capita water withdrawal across the entire U.S. The predictive model based on the method of random forest was selected as the best model, since it out-performed all other statistical models in-terms of both goodness-of-fit and out-of-sample predictive accuracy.

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Our results identified irrigated farming - especially in the states such as Nebraska and Arkansas – and coal mining especially in states such as Wyoming, West Virginia and Kentucky as the most water-intensive anthropogenic activities. Even though mining withdrawals constitute a small fraction of the overall water use in the U.S., its share has increased by 40% since 2005 (Maupin et al., 2014).

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The water intensity of thermoelectric generation was less than initially hypothesized. According to the USGS, the reduced water withdrawals for thermoelectric power generation over the years can be attributed to a reduction in coal consumption and increased use of natural gas, as well as the newer power plants being equipped with more water-efficient cooling technologies. The USGS also reports declined industrial water with-

486 drawals due to higher efficiencies in industrial activities and an emerging emphasis on
487 water reuse and recycling in industrial processes (Maupin et al., 2014).

488 Climatic conditions such as precipitation and heating-degree days were also found
489 to be important predictors of per-capita water withdrawal. Drier conditions (i.e., total
490 annual precipitation less than 600) were intuitively found to be associated with higher
491 water withdrawals. However, counter-intuitively, we found colder conditions i.e., HDD
492 > 3000 which is mostly observed in the North-Central parts of the U.S., such as North
493 Dakota, Minnesota, Wyoming and Montana – to be associated with higher water use.
494 This higher water use might be attributed to hydraulic fracturing for oil and gas and other
495 mining activities beyond coal mining in these states. While the total, per-capita water
496 withdrawals are lower in more urbanized states, the water withdrawal in the public sup-
497 ply is positively associated with urbanization.

498 Using the developed predictive model, we were able to infer the first-order sensi-
499 tivity of the projected changes in the water withdrawal to changing climate conditions
500 such as precipitation. Our analysis results revealed a distinct north-south gradient in the
501 projected changes of the water withdrawal pattern across U.S. (mostly between $\pm 10\%$),
502 with the southern (northern) states showing projected increase (decrease) in future wa-
503 ter usages in response to the projected changes in mean annual precipitation by the end
504 of Century under the RCP8.5 scenario. In a similar fashion, our data-driven modeling
505 framework allows for analyzing and documenting the sensitivity of future changes in wa-
506 ter withdrawal in response to other climatic (e.g., HDD changes) and socioeconomic fac-
507 tors (e.g., changes in farmland expansion, urbanization, energy generation); either in-
508 dividually (considering one at a time) or in combination.

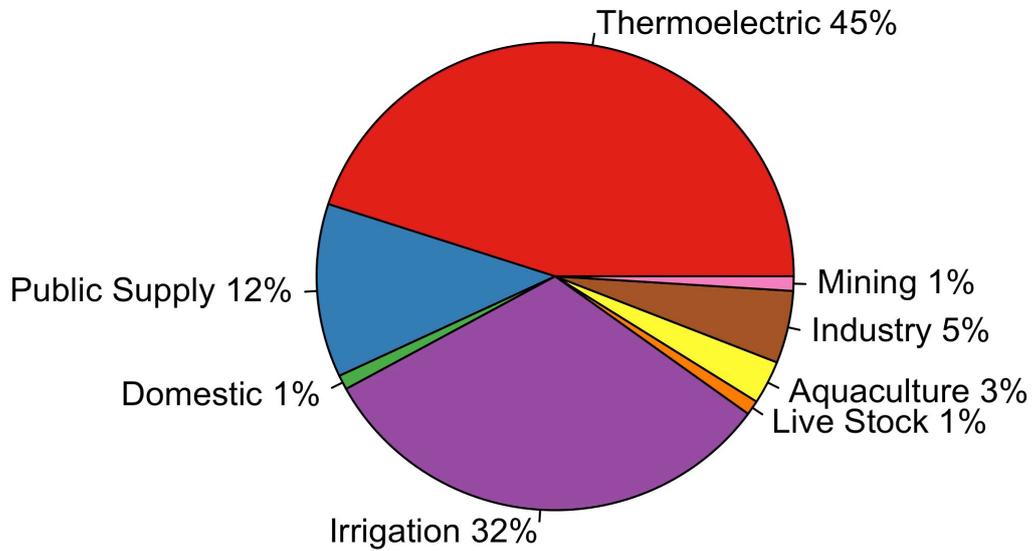
Table 1. Summary of models performance given as correlation coefficient (R^2), fitted Root Mean Square Error (RMSE; million gal/day/person), and Leave one out cross validation (LOOCV) RMSE. Each model is trained and tested using all available data records for the period 1991-2010.

Model	R^2	RMSE	LOOCV RMSE
Mean-ONLY	–	2.60	2.62
Multiple Linear Regression (MLR)	0.57	1.71	1.84
Generalized Additive Model (GAM)	0.61	1.62	1.62
Multivariate Adaptive Regression Splines (MARS)	0.85	0.99	1.40
Random Forest (RF)	0.97	0.47	0.98

Table 2. Summary of models predictive accuracy. Each Model is trained using 1991-2005 data and tested using 2006-2010 data. Summary performance is presented here in terms of correlation coefficient (R^2), fitted Root Mean Square Error (RMSE; million gal/day/person), Leave one out cross validation (LOOCV) RMSE, and prediction RMSE (for the test data). See Appendix D for more details on LOOCV-RMSE.

Model	R^2	RMSE	LOOCV RMSE	Prediction RMSE
Mean-ONLY	–	2.75	2.77	2.11
Multiple Linear Regression (MLR)	0.59	1.76	2.00	1.52
Generalized Additive Model (GAM)	0.65	1.63	1.68	1.31
Multivariate Adaptive Regression Splines (MARS)	0.95	0.60	1.57	1.35
Random Forest (RF)	0.97	0.48	1.00	0.79

Pie Chart of Water Withdrawal Breakdown for 2010



Map of Total Water Withdrawal PerCapita (million gallons/day/person)

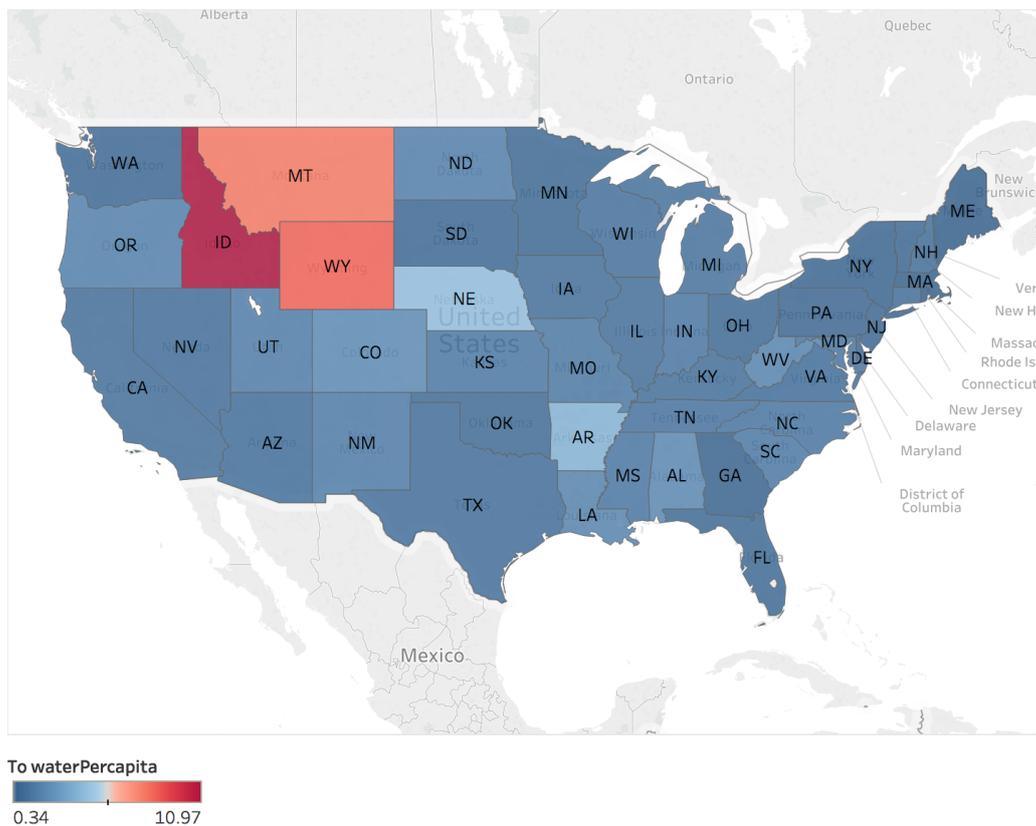


Figure 1. Top: The breakdown of US-wide water withdrawals across the eight major sectors during the period 2006-2010. Bottom: Spatial distribution of the U.S. wide per-capita water withdrawal (in million gallons per-day).

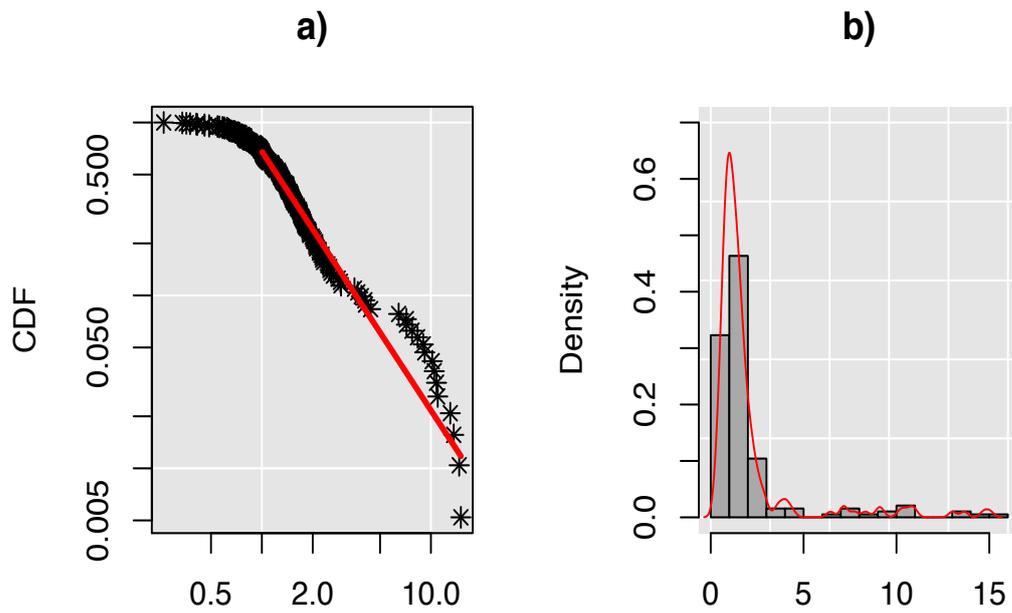


Figure 2. The empirical distribution of per-capita water withdrawals (in million gallons per day) for the period 1991-2010; (a) the red line shows that power-law fits the tail of the empirical cumulative distribution reasonably well (b) the histogram of per-capita water demand with overlain kernel density line (in red).

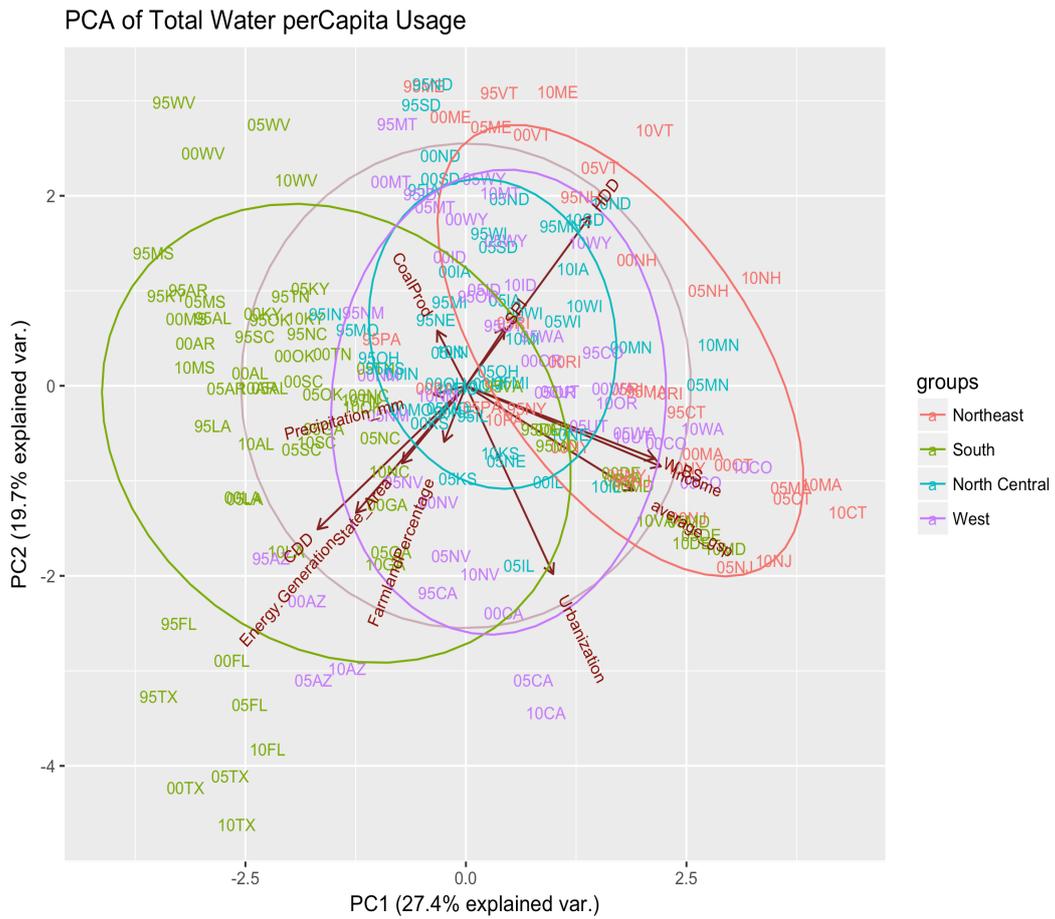


Figure 3. Principal Component Analysis (PCA) biplot of the per-capita water usage (in million gallons per-day) for the period 1995-2010. The states are color-coded based on their proximity to water bodies and the two digits next to the state codes indicate the year associated with the water use data for the state.

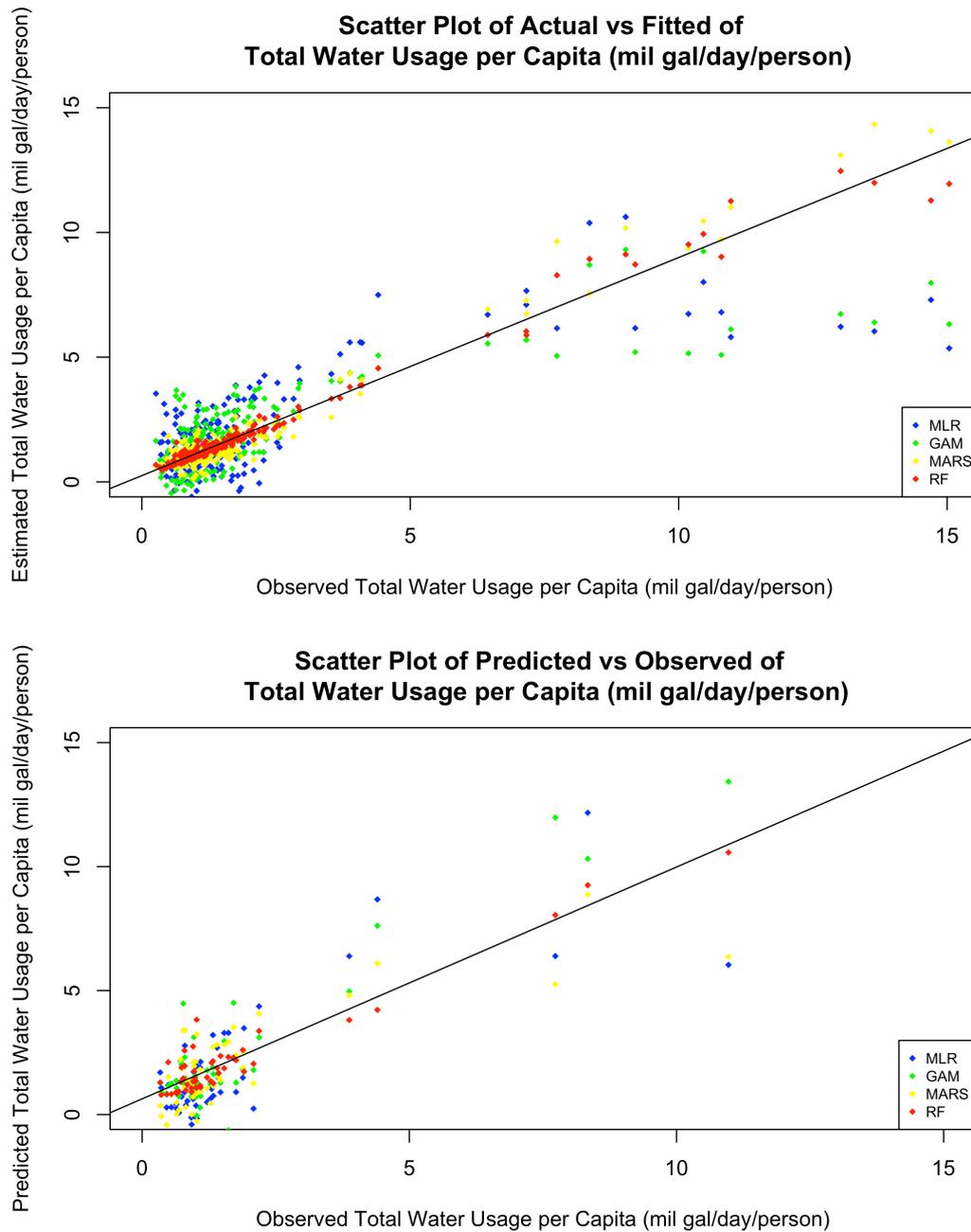


Figure 4. Top: Scatter plot of observed versus estimated values of per-capita water withdrawal (in million gallons per-day) using data of 1995-2010. Bottom: Scatter plot of observed versus predicted values of per-capita water usage (in million gallons per-day) using data of 2006-2010. In the latter case, the models were trained using data of 1995-2005, and the testing was conducted in an independent period of 2006-2010.

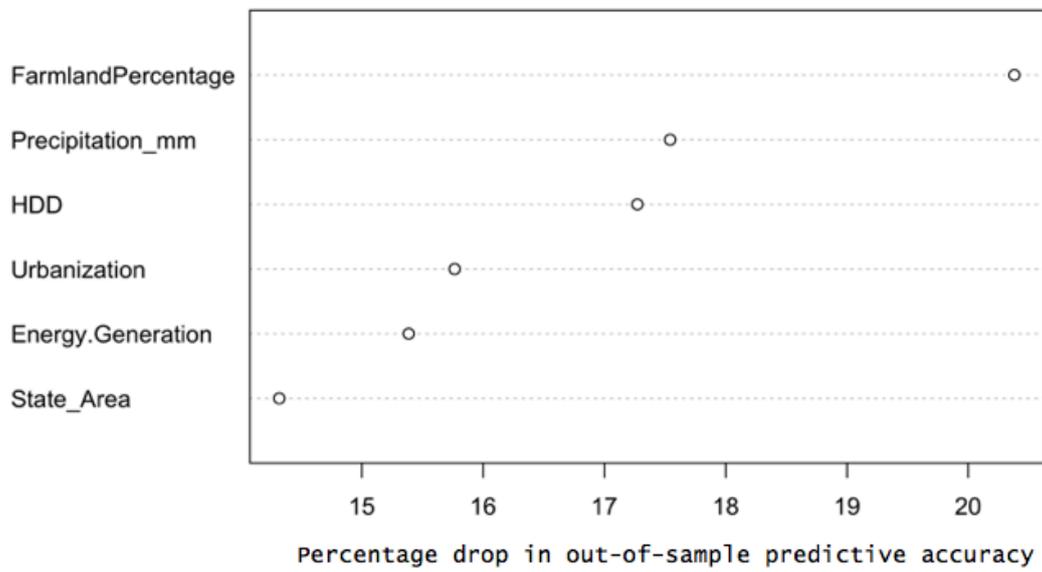


Figure 5. List of the most important predictors identified for the per-capita water withdrawal predictions, presented here as the percentage drop in the predictive accuracy for the out-of-sample datasets. The selected predictors are ranked from the most to least influential ones (top to bottom).

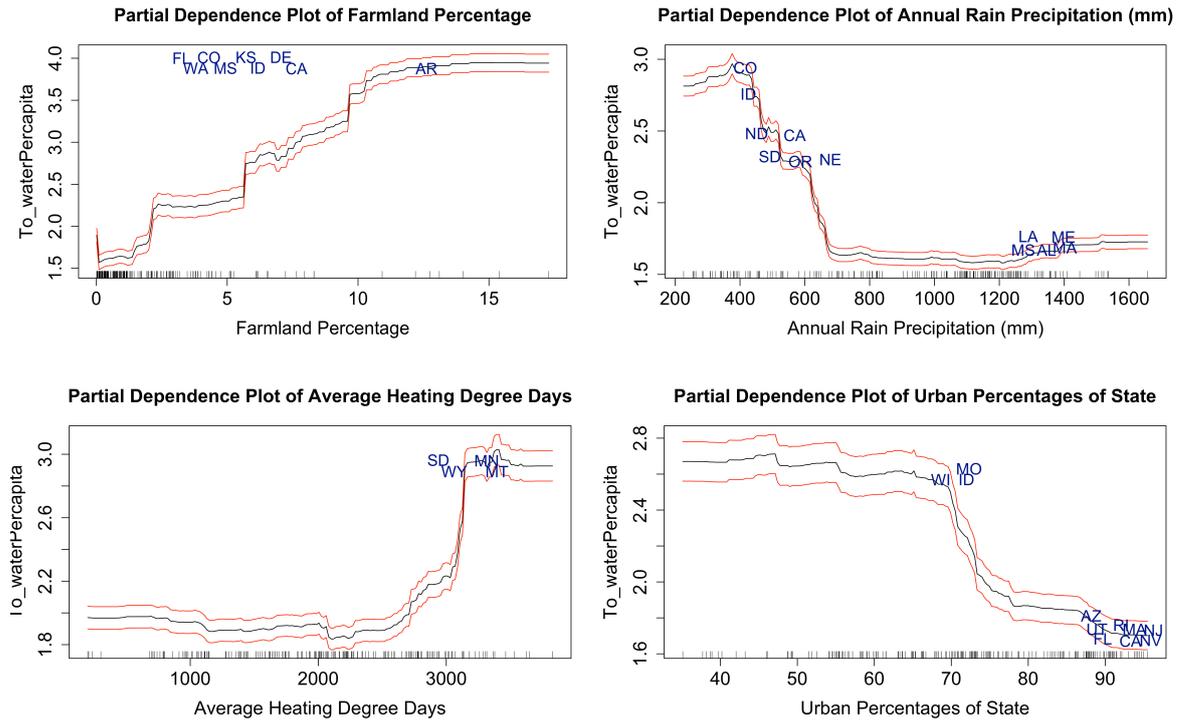


Figure 6. Partial dependency plot (PDP) for the fraction of irrigated farmland, annual precipitation, average heating degree days, and percentage of urban areas; depicting their sensitivity on the per-capita water withdrawal (in million gallons per-day). The two letters on the plot corresponds to the states, the black line the mean values, and the red lines the 95% confidence intervals.

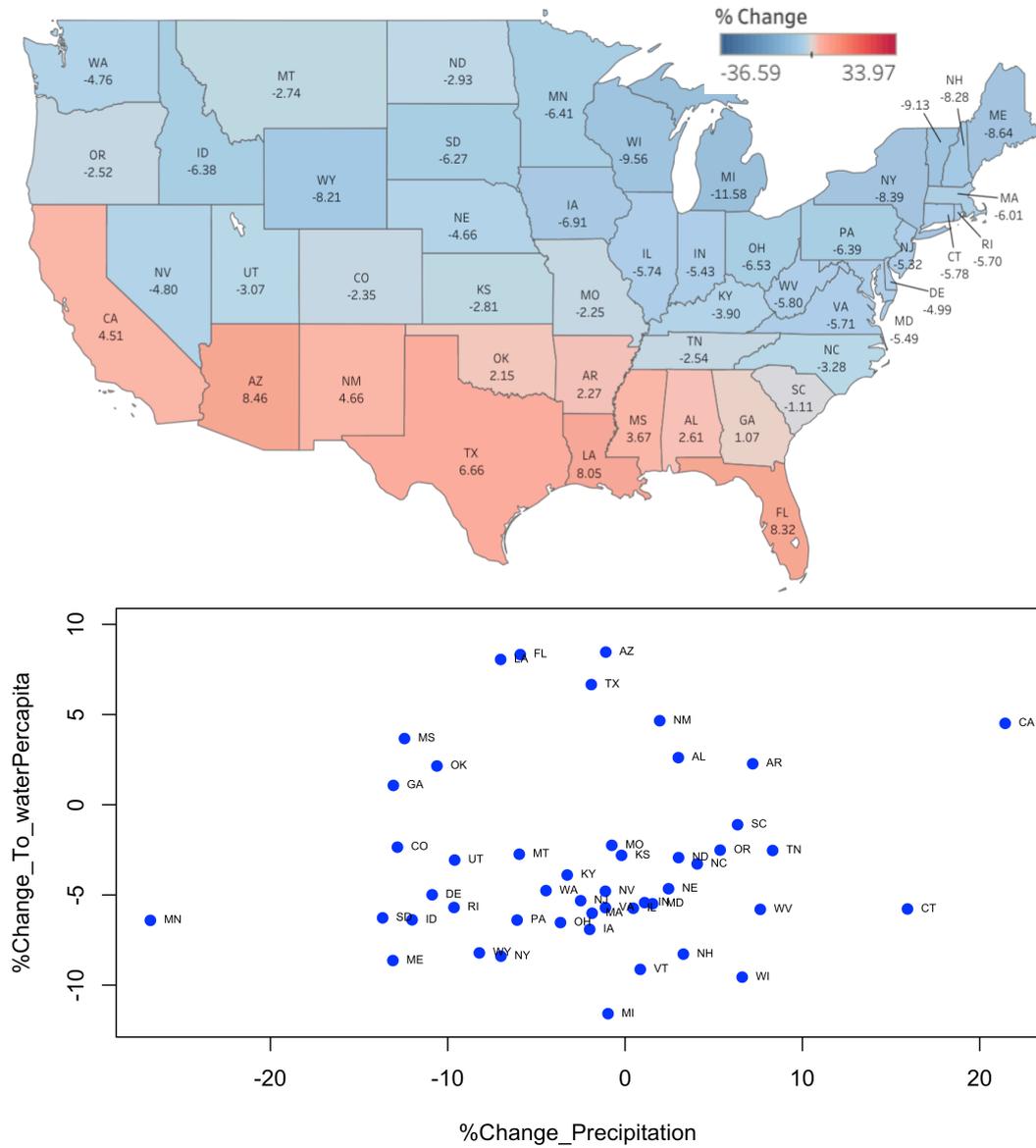


Figure 7. Spatial distribution of ensemble mean changes in the per-capita water withdrawal patterns across the U.S. in response to the future precipitation changes. Ensemble means were estimated based on the modeled WW values using the mean annual precipitation estimates from the five CMIP5 GCMs, while other predictors were kept constant at nominal values (see the corresponding texts for more details). Changes in the WW estimates corresponds to the future period (2070-2085) under the RCP8.5 scenario, relative to the reference estimates of the contemporary conditions (1995-2010). Bottom panel shows the scatter plot of percentage changes between precipitation and total per-capita water withdrawal.

509 **A Generalized Linear Model (GLM)**

510 GLM is the extension of Ordinary Linear Regression (OLR). GLM still retains all
 511 the assumption of OLR, it allows predictors to be categorical and allows interactions be-
 512 tween predictors. The simplest form of GLM is define as one of the most widely used
 513 methods for function approximation. GLM can be mathematically summarized as be-
 514 low:

$$515 \quad y_i = \beta_0 + \beta_1 x_{i,1} + \beta_2 x_{i,2} + \dots + \beta_n x_{i,n} + \epsilon_i \quad (\text{A.1})$$

516 where the stochastic error is assumed to be normally distributed as: for all $\epsilon_i \sim$
 517 $N(0, \sigma^2)$. Each β_j describes the slope of predictor $x_{(i,j)}$. Sometimes transformation of
 518 original variables (such as polynomials) are used to improve the performance of the mod-
 519 els. Multiple linear Regression (MLR) is popular because they can be easily fitted (even
 520 with limited data) and they are easily interpretable. However, their ‘rigid’ structure of-
 521 ten fail to approximate the true function, especially when response is a complex (non-
 522 linear) function of input variables. Their predictive accuracy is therefore often inferior
 523 to more flexible models (James, Witten, Hastie, & Tibshirani, 2013).

524 **B Generalized Additive Models (GAM)**

525 GAM is a natural extension from GLM, in order to preserve the additive model
 526 while extending to nonlinear relationship between the response and predictors (Hastie
 527 et al., 2009). GAM is a non-parametric (local parametric) fitting procedure where the
 528 conditional expectation of y is related to the input variables space as shown below:

$$529 \quad y_i = \beta_0 + \sum_{j=1}^p f_j(x_{i,j}) + \epsilon_i \quad (\text{B.1})$$

530 where $f_j(x_{i,j})$ is a smoothing splines over the p -dimensional input space, with the
 531 number of observations running from $i = 1, \dots, n$. GAM relaxes the linearity assump-
 532 tion of multiple linear regression with smoothing functions $f_j(x_{i,j})$. This allows for cap-
 533 turing the non-linear relationship between the predictors and the response variable. The
 534 flexibility of generalized additive model often result in better approximating the true func-
 535 tion and therefore often outperform GAM in predictive accuracy.

536 **C Multivariate Adaptive Regression Splines (MARS)**

537 MARS is a non-parametric regression techniques developed by Friedman (1991).
 538 It extends the use of piecewise linear basis function of form $(x-t)_+$ and $(x-t)_-$, where

$$539 \quad (x-t)_+ = \begin{cases} x-t & x > t \\ 0 & otherwise \end{cases} \quad (\text{C.1})$$

$$540 \quad (x-t)_- = \begin{cases} t-x & x < t \\ 0 & otherwise \end{cases} \quad (\text{C.2})$$

541 And MARS has the function form of

$$542 \quad f(X) = \beta_0 + \sum_{m=1}^M \beta_m h_m(X) \quad (\text{C.3})$$

543 where each $h_m(X)$ is a function in form of piecewise linear basis function, or the
 544 product of two or more such functions. The coefficients β_m are estimated by minimiz-
 545 ing the residual sum of squares given the choices of $h_m(X)$ (Hastie et al., 2009).

546 D Bias-variance trade-off

547 Predictive performance of a statistical model depends on its capability to yield ac-
 548 curate predictions for an independent test sample. Generally simple models are more sta-
 549 ble, but do not adequately estimate the structure of the true function – and therefore
 550 are high in bias. Complex models can approximate the shape of the true function, more
 551 effectively, but they are prone to over-fitting – and therefore have high variance. Bias-
 552 variance trade-off lies at the heart of developing models with high generalization power
 553 add references. Cross-validation is one of the most widely used methods in balancing bias
 554 and variance. We use the leave-one-out cross validation (LOOCV) to estimate predic-
 555 tive accuracy. The LOOCV procedure is defined as holding out one data as a test data
 556 and use the rest of the training data. Model generated from the training data is the used
 557 to predict the test data and we will calculate the MSE of that point. LOOCV MSE is
 558 defined by

$$559 \quad LOOCVMSE = \frac{1}{n} \sum_{i=1}^n (y_i - \hat{y}_i)^2 \quad (D.1)$$

560 where i represents the iteration of one data left out, y_i represents the true value
 561 of the i^{th} iteration, \hat{y}_i represents the predicted value and n the length of data.

562 Acronyms

563 **BAU** Business as Usual
 564 **CDD** Cooling Degree Days (°F)
 565 **CPC** Climate Prediction Center
 566 **CMIP5** Coupled Model Intercomparison Project (Phase 5)
 567 **EIA** Energy Information Association
 568 **EPA** Environmental Protection Agency
 569 **EPRI** Electric Power Research Institute
 570 **GAM** Generalized Additive Model
 571 **GCM** Global Circulation Model
 572 **GDP** Gross Domestic product
 573 **GFDL-ESM2** Geophysical Fluid Dynamics Laboratory-Earth System Models
 574 **GLM** Generalized Linear Model
 575 **GSP** Gross State Product (millions of USD measured in 2009 real dollars)
 576 **HadGEM2-ES** Met Office Hadley Centre Model-Earth System
 577 **HDD** Heating Degree Days (°F)
 578 **IPSL-CM5A-LR** Institut Pierre Simon Laplace Model-5 Component models
 579 **ISI-MIP** Inter-Sectoral Impact Model Intercomparison Project
 580 **NEMS** National Energy Modeling Systems
 581 **NOAA** National Oceanic and Atmospheric Administration
 582 **MARS** Multivariate Adaptive Regression Splines
 583 **MIROC-ESM-CHEM** Model for Interdisciplinary Research on Climate-Earth Sys-
 584 tem Models
 585 **NorESM1-M** Norwegian Earth System Model 1 - Medium resolution
 586 **PDP** Partial Dependence Plot
 587 **RCP** Representative Concentration Pathway

588 **RF** Random Forest
 589 **SPI** Standardized Prediction Index
 590 **U.S.** United States
 591 **USD** United States Dollar (\$)
 592 **USGS** United States Geological Survey

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 599 edge the World Climate Research Programme’s Working Group on Coupled Modelling,
 600 which is responsible for CMIP, and we thank the climate modeling groups for produc-
 601 ing and making available their model output. All data-sets used in this study are col-
 602 lected from publicly available sources as detailed in Section 3. The source codes of the
 603 different statistical algorithms used in this study can be acquired from the correspond-
 604 ing authors, and they will be hosted on the on-line repository platform (Github).

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