Data-driven Estimation of Groundwater Level Time-Series Using Comparative Regional Analysis

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Abstract

A new method is presented to efficiently estimate daily groundwater level time series at unmonitored sites by linking groundwater dynamics to local hydrogeological system controls. The presented approach is based on the concept of comparative regional analysis, an approach widely used in surface water hydrology, but uncommon in hydrogeology. The method uses regression analysis to estimate cumulative frequency distributions of groundwater levels (groundwater head duration curves (HDC)) at unmonitored locations using physiographic and climatic site descriptors. The HDC is then used to construct a groundwater hydrograph using time series from distance-weighted neighboring monitored (donor) locations. For estimating times series at unmonitored sites, in essence, spatio-temporal interpolation, stepwise multiple linear regression, extreme gradient boosting, and nearest neighbors are compared. The methods were applied to ten-year daily groundwater level time series at 157 sites in alluvial unconfined aquifers in Southern Germany. Models of HDCs were physically plausible and showed that physiographic and climatic controls on groundwater level fluctuations are nonlinear and dynamic, varying in significance from "wet" to "dry" aquifer conditions. Extreme gradient boosting yielded a significantly higher predictive skill than nearest neighbor and multiple linear regression. However, donor site selection is of key importance. The study presents a novel approach for regionalization and infilling of groundwater level time series that also aids conceptual understanding of controls on groundwater dynamics, both central tasks for water resources managers.

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8					
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10	Key Points:				
11	• Presents method for estimation of daily groundwater levels through transfer of head				
12	duration curves based on similarity of site characteristics at monitored sites.				
13	• Nonlinearity of controls on groundwater levels favors use of Machine Learning (e.g.,				
14	regression trees) over multiple linear regression for prediction.				
15	• Investigates the dynamic nature of controls on groundwater levels, which is central for				
16	studies of recharge seasonality, droughts and floods.				
17					

18 Abstract

19 A new method is presented to efficiently estimate daily groundwater level time series at 20 unmonitored sites by linking groundwater dynamics to local hydrogeological system controls. The 21 presented approach is based on the concept of comparative regional analysis, an approach widely 22 used in surface water hydrology, but uncommon in hydrogeology. The method uses regression 23 analysis to estimate cumulative frequency distributions of groundwater levels (groundwater head 24 duration curves (HDC)) at unmonitored locations using physiographic and climatic site 25 descriptors. The HDC is then used to construct a groundwater hydrograph using time series from distance-weighted neighboring monitored (donor) locations. For estimating times series at 26 27 unmonitored sites, in essence, spatio-temporal interpolation, stepwise multiple linear regression, 28 extreme gradient boosting, and nearest neighbors are compared. The methods were applied to ten-29 year daily groundwater level time series at 157 sites in alluvial unconfined aquifers in Southern 30 Germany. Models of HDCs were physically plausible and showed that physiographic and climatic 31 controls on groundwater level fluctuations are nonlinear and dynamic, varying in significance from 32 "wet" to "dry" aquifer conditions. Extreme gradient boosting yielded a significantly higher 33 predictive skill than nearest neighbor and multiple linear regression. However, donor site selection 34 is of key importance. The study presents a novel approach for regionalization and infilling of 35 groundwater level time series that also aids conceptual understanding of controls on groundwater 36 dynamics, both central tasks for water resources managers.

37 1 Introduction

Groundwater head observations are the basis for most investigations in hydrogeology. However, boreholes for groundwater observation as well as corresponding groundwater level time series are often scarce and unevenly distributed in both space and time. This is a disadvantage for effective management of groundwater resources at the regional scale (Butler et al., 2021), where water managers assess the current and future status of groundwater resources (Lóaiciga & Leipnik, 2001). In consequence, methods are needed to estimate groundwater head time series at ungauged sites.

45 Two main approaches are commonly used by hydrogeologists to predict temporal changes 46 in groundwater head at a given site, (a) numerical and (b) statistical models. The typical approach 47 is to implement a process-based, numerical groundwater flow model. However, numerical models 48 typically require large amounts of data and effort, while investigators commonly are confronted 49 with a lack of comprehensive description and documentation of the subsurface. This results in 50 significant uncertainty, both regarding conceptualization and parametrization (e.g. Enemark et al., 51 2019). Dealing with this uncertainty leads to a tedious and time-consuming process to construct, 52 calibrate, and run these process-based models (Bakker & Schaars, 2019). Additionally, models for 53 meaningful local projections at large spatial scales are not yet available (Berg & Sudicky, 2019). 54 An alternative to regional scale modelling with less need for detailed subsurface description are 55 lumped (rainfall-runoff) hydrological models with a groundwater component (Barthel & Banzhaf, 56 2016). However, these models are problematic as they usually imply oversimplification of the 57 groundwater component, disregarding the local descriptors of hydrogeological systems and their 58 3-dimensional setup (Barthel & Banzhaf, 2016; Butler et al., 2021). Generally, lumped models 59 may provide adequate descriptions of groundwater systems only for simple hydrogeological 60 situations such as shallow, unconfined aquifers, but not for more complex systems, such as deep and confined aquifers. 61

62 A different type of approach requiring only measured groundwater level data for groundwater time series estimation are parametric or data-driven methods. This approach requires 63 64 few data on local system descriptors, while often long and measurement-dense series of input 65 signal and groundwater measurements are necessary to achieve good calibrations. In contrast to 66 groundwater-gradient driven methods, data-driven methods either use spatio-temporal 67 geostatistics (e.g. Ruybal et al., 2019; Varouchakis et al., 2022) or transfer net precipitation input 68 into groundwater level changes (Z. Chen et al. (2002)). However, available methods predict 69 groundwater level only at monthly or annual resolution and consequently do not capture the large 70 intra-annual and intra-monthly variability of groundwater dynamics (e.g. Heudorfer et al., 2019). 71 An approach to predict time series at higher temporal scales are transfer functions, that can be used 72 to yearly, monthly and daily temporal resolutions, such as impulse-response functions (e.g. 73 Collenteur et al., 2019; Marchant & Bloomfield, 2018; Von Asmuth, 2012) or artificial neural 74 networks (c.f. Rajaee et al., 2019; Wunsch et al., 2022). However, no formal method is known to 75 transfer information from such models from monitored to unmonitored aquifers, although recently 76 attempted in streamflow (Kratzert et al., 2019). This means that these methods can only make 77 predictions when sufficient local time series data are available (e.g., 10 years weekly data (Wunsch 78 et al., 2021)).

79 In summary, neither numerical models nor the currently available data-driven tools provide 80 a straightforward approach to estimate daily groundwater levels at unmonitored sites to aid 81 regional scale management. Therefore, new and complementary methodologies are required to 82 overcome scarcity and patchy data distribution. Such approaches should be less data hungry than 83 numerical models, yet account for local hydrogeological conditions and allow prediction at high 84 temporal resolution despite limited local data availability. In surface-water-orientated hydrology, 85 data scarcity has been countered with approaches of classification and similarity analysis, 86 embraced by the hydrological community particularly within the PUB initiative (Predictions in 87 Ungauged Basins; (Blöschl et al., 2013; Hrachowitz et al., 2013; McDonnell & Woods, 2004; 88 Sivakumar & Singh, 2012; Wagener et al., 2007). These concepts attempt to systematically link 89 the physical form and structure of catchments to their functioning by comparative analysis. Such 90 links can then be used to transfer information to similar systems for prediction, i.e., regionalization 91 or spatio-temporal interpolation. However, such approaches are rarely considered in groundwater 92 research, which is pointed out by various authors, e.g., Barthel et al. (2021); de Marsily et al. 93 (2005); Green et al. (2011); Voss (2005). Recently, a number of studies initiated the 94 implementation of these approaches in groundwater, quantitatively connecting groundwater 95 response to physiographic and climatic descriptors (Boutt, 2017; Giese et al., 2020; Haaf & 96 Barthel, 2018; E. Haaf et al., 2020; Heudorfer et al., 2019; M. Rinderer et al., 2017; M. Rinderer 97 et al., 2019; M. Rinderer et al., 2014; Michael Rinderer et al., 2016). These approaches, however, 98 have not yet been exploited to predict daily groundwater levels at unmonitored sites.

99 When looking for methodological inspiration in the body of literature within the surface 100 water community, and more specifically the PUB initiative, a large majority of approaches use 101 regionalization mainly as a tool to calibrate lumped rainfall-runoff models at unmonitored sites 102 (He et al., 2011; Hrachowitz et al., 2013). As mentioned above, such lumped models are often not 103 useful for describing groundwater dynamics and, when available, are time-consuming to set up 104 and calibrate (Jackson et al., 2016; Mackay et al., 2014). Simpler statistical methods for 105 regionalization of streamflow time series, however, have been proposed by e.g. Shu and Ouarda 106 (2012) based on Hughes and Smakhtin (1996). These methods make use of the characteristic 107 relationship between flow duration curve (FDC; cumulative frequency of time where a flow is 108 equaled or exceeded) and physiographic and climatic site descriptors, a relationship that is well 109 investigated (Yokoo & Sivapalan, 2011). FDCs in surface water hydrology are commonly used to

110 study the flow regime throughout the range of discharges and integrate effects of climate, 111 topography, geology, and also anthropogenic activity (Ridolfi et al., 2020; Sugiyama et al., 2003; 112 Vogel & Fennessey, 1995). This implies that the shape of a specific FDC is theoretically inferable from site descriptors. The technique evaluated in this study takes advantage of this through 113 114 estimation of duration curves at unmonitored (target) sites based on similarity to neighboring donor 115 sites. Then, from the estimated duration curve, time series are reconstructed at the target site into 116 a daily time series (Hughes & Smakhtin, 1996; Mohamoud, 2010; Shu & Ouarda, 2012; Smakhtin, 117 1999).

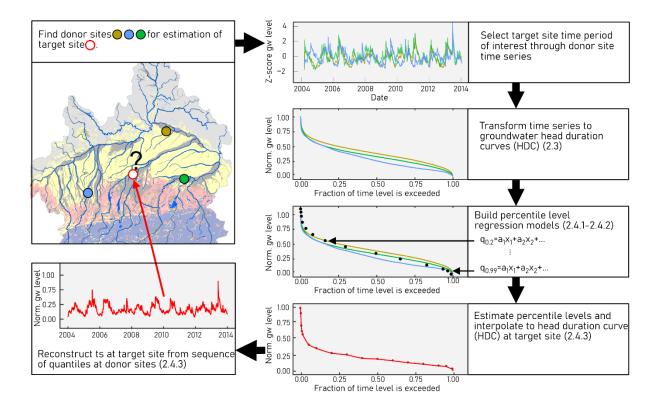
118 Cumulative frequency or duration curves of groundwater heads are not as broadly used for 119 studying groundwater resources, except when for example analyzing the relative state of 120 groundwater storage (e.g. Maxe, 2013). Giese et al. (2020) estimated aggregates (indices) of head 121 duration curves (HDC) and linked differences in shapes to local, intermediate, and regional 122 groundwater flow patterns. Ezra Haaf et al. (2020) found correlation between HDC indices and 123 map-derivable physiographic and climatic site descriptors. These are indications that alike 124 streamflow, system controls are integrated in groundwater level regimes and may be exploited by 125 analysis of duration curves.

126 Accordingly, regionalization and subsequent estimation of daily time series at unmonitored 127 sites through duration curves of groundwater head is evaluated in this paper. Hereby the approach 128 is based on the methodology proposed by Shu and Ouarda (2012) for streamflow. It is adapted to 129 groundwater, where groundwater head duration curves as well as groundwater-relevant and map-130 derivable site descriptors are used. Within surface-water, this method has only been tested using 131 stepwise multiple linear regression (MLR). In this study, a comparison is carried out with 132 estimation through averaging of the nearest neighbor sites (NN), MLR, and extreme gradient 133 boosting (XGB). XGB can represent nonlinear relationships between groundwater dynamics and 134 site descriptors and has shown to be powerful in e.g., recharge studies (Naghibi et al., 2020). In summary, a method is evaluated that may be used when aquifer and time series data at a site of 135 interest are unmonitored. The regionalization approach is applied to unconfined, alluvial aquifers 136 137 in a humid climate in Southern Germany at unmonitored sites using solely map-derivable site 138 descriptors and data from neighboring locations.

139 **2 Method and Data**

140 **2.1 General strategy**

141 The methodology of estimating groundwater level time series at an unmonitored site, is 142 based on information from donor sites and requires the steps as explained in Figure 1. In the 143 beginning, donor sites are selected with a time series period that is of interest for target site 144 estimation. Next, time series are transformed to HDCs, and at 15 fixed percentile levels, models 145 are constructed based on multiple regression analysis and gradient boosted regression trees, and 146 logarithmically inter- and extrapolated (section 2.4.1-2.4.2). Finally, time series at ungauged sites 147 are then reconstructed with a distance-based weighting method using the sequence of records from 148 donor sites (section 2.4.3). For performance comparison, time series are also evaluated using only 149 a distance-based average of time series from donor sites, further called Nearest-Neighbour (NN). 150 Then, the number of neighbors and the performance of daily groundwater level estimations at 151 target sites are evaluated using leave-one-out cross-validation (2.5). The models that are used for 152 estimation of time series are then checked for plausibility (section 2.6). In section 2.7 the case data 153 set is described, which is further analyzed using cluster analysis to understand results with regard 154 to different groundwater regimes and systems. All data analysis was carried out by using the 155 programming language R (R Development Core Team, 2022).



156

157 Figure 1. Principle steps to estimate groundwater level time series at unmonitored sites using
158 the head duration curve methodology.

159 2.2 Data Selection and Processing

160 Groundwater level time series are selected from a data set described by E. Haaf et al. (2020). The data set contains groundwater level time series from the Upper Danube catchment in 161 162 Bavaria, Southern Germany, with available geological information and absence of patterns of 163 direct anthropogenic impact (for a more detailed explanation refer to Heudorfer et al. (2019)). From this data set observation wells were selected that come (1) with continuous daily time series 164 and at least 10 year record length, (2) less than 1% missing data, which are (3) concurrent with a 165 166 record period 2004–2014. The resulting set of 157 obervation wells are mostly located in shallow, quaternary sediments in river valleys and fluvial sand as well as in gravel deposits, with a few 167 168 boreholes located in deeper tertiary sediments. All wells are classified as penetrating unconfined 169 aquifers. Then, at each site, 47 physiographical and meteorological descriptors were derived, 170 described in detail in Ezra Haaf et al. (2020). In addition to Ezra Haaf et al. (2020), percentage of 171 land cover within a 3 km radius of each site was derived from the CORINE land cover data set 172 (Bossard et al., 2000). Table 1 shows selected descriptors that are most important for models on this study and therefore discussed in more detail. Remaining descriptors can be found in the supporting information SI (Table S1). Descriptors are called predictors when in context of regression models.

- 176
- 177 Table 1. Descriptive statistics of physiographic and climatic descriptors, discussed in the
- 178 paper. Class of variable in parenthesis: (G) Geology, (M) Morphology, (L) Land cover, (B)
- 179 Boundaries and (C) Climate.

Variable	Description	Range		Unit
		Minimum	Maximum	-
dist_stream (B) ‡	Estimated distance from well to nearest stream (main rivers)	6	10958	m
well_elevation (B)	Estimated Elevation of well	310	839	m asl.
P_avg (C)	Mean annual precipitation	675	1613	mm
T_avg (C)	Mean annual temperature	6.4	9.3	°C
SI (C)	Seasonality index of precipitation	.11	.31	-
A_thickness (G)	Average thickness of saturated zone	1	50.1	m
A_Depth (G)	Bottom of formation	3	110	m
Depth_to_GW (G)	Average depth to Water table	0.3	39.8	m
Broadleaved_forest (L)	% of 3 km buffer occupied by broadleaved forest	0	44.5	%
Coniferous_forest (L)	% of 3 km buffer occupied by coniferous forest	0	93.5	%
Urban (L)	% of 3 km buffer occupied by urban fabric	0	74.9	%
slp_sk (M) †	Mean slope	0/-0.1	1.95/2.6	-
twi (M)	Mean value of Topographic Wetness index	5.8	8.9	-

180 † skewness was calculated for local and regional scale respectively. For these, the ranges are given seperated by a 181 slash l/r.

182 **2.3 Transformation to head duration curves (HDCs)**

In a first step, groundwater head time series were normalized. Subsequently, duration curves of groundwater levels were calculated at each site. This was done, by first ranking all *n* observed, normalized (on a 0-1 scale) groundwater levels l_i , i = 1, 2, ..., n in descending order, where *i* is the rank of an observation. The head duration curve (HDC) is then constructed following the Weibull plotting formula (Sugiyama et al., 2003):

188
$$p_i = P(L \ge l_i) = \frac{i}{n+1},$$
 (1)

189 where p_i is the percentage of time that a given level l_i is equaled or exceeded. Groundwater 190 level or head duration curves are subsequently the plot of percentage level p_i against the 191 corresponding level l_i (as seen in Figure 1).

192 2.4 Regression analysis for percentile models

To be able to estimate the duration curve at an ungauged site, forward stepwise regression (MLR, see section 2.4.1) and extreme gradient boosting (XGB, see section 2.4.2) were applied to build models from physiographic and climatic predictors at selected percentage level (0.1%, 0.5%, 1%, 5%, 10%, 20%, 30%, 40%, 50%, 60%, 70%, 80%, 90%, 95%, 99%). Models are fit using a nested cross-validation approach on 80% training data with 20% hold-out data on which evaluation is performed. Models were trained 30 times by leaving out one group each time and then evaluating against predictions in the left-out group.

200 2.4.1 Construction of percentile models with MLR

201 Multiple linear regression models at selected percentage levels are built using a selective 202 inference framework. Selective inference adjusts p-values for the effect of sequentional selection 203 of variables (Taylor & Tibshirani, 2015). This is necessary since conventional stepwise regression 204 leads to an overestimation of the strength of apparent relations. The consequence of conventional 205 models is therefore selection of non-significant predictors and therefore overfitting (Taylor & 206 Tibshirani, 2015). Instead of using p-values based on the t-test for forward selection, the procedure 207 is here stopped based on the false discovery rate (exceeding 0.1; (G'Sell et al., 2016). The selected 208 variables are then used to build a regression relationship for the training data set with n209 observations (from well locations) and percentage levels, $p = 1, 2 \dots n$, where H_p is the percentile of the normalized head H and x_p the selected climatic and physiographic descriptors with the 210 211 following form:

212
$$H_p = \beta_0 + \sum_j x_{pj} \beta_j + \epsilon_p, \qquad (2)$$

213 errors ϵ_p being independent and normally distributed and where β is a vector of model 214 parameters that are estimated.

215 2.4.2 Construction of percentile models with XGB

216 Alternative models for each percentile were constructed using extreme gradient boosting, 217 an implementation of boosted regression trees (Friedman, 2001). Hereby, the xgb.train function from the XGBoost R package (T. Chen & Guestrin, 2016) was used to predict H_p based on the 218 219 entire set of climatic and physiographic descriptors. To optimize the model fit but reduce risk of 220 overfitting, two further steps were carried out, after the 80/20 hold-out split mentioned above. 221 After this, hyperparameters were determined on the training data using 5-fold cross validation, 222 using the performance measure root mean square error (RMSE). Finally, after cross-validation, the 223 risk for overfitting was reduced by stopping the ensemble at the number of decision trees, where 224 the difference between training and evaluation error reaches a minimum.

225 **2.4.3** From percentile models to estimated time series

Once percentile levels are predicted for a given target site using XGB and MLR models, logarithmic interpolation is used to estimate percentiles of groundwater levels between the percentage points in order to construct the entire duration curve. The percentile to be estimated is found by identifying the closest (modelled) fixed percentage levels p_i above and p_{i-1} below and their corresponding groundwater heads H_i and H_{i-1} . The groundwater level H can then be found using the following equation:

232
$$\ln(H) = \ln(H_i) + \frac{\ln(H_{i-1}) - \ln(H_i)}{p_{i-1} - p_i} \times (p - p_i)$$
(3)

In cases where percentiles are estimated that are larger than the highest percentage point or lower than the lowest (modelled) percentage point, logarithmic extrapolation is used. Hereby, the closest two percentage points are found (p_{n1}, p_{n2}) and the corresponding groundwater levels (H_{n1}, H_{n2}) . Extrapolating to the percentile *p* is done using the equation below.

237
$$\ln(H) = \ln(H_{n1}) + \frac{\ln(H_{n1}) - \ln(H_{n2})}{p_{n1} - p_{n2}} \times (p - p_{n2})$$
(4)

Reconstruction of the groundwater level time series from interpolated duration curves can then be carried out following the principle given by Smakhtin (1999) for streamflow estimation. Groundwater levels H_t at the target site are estimated by looking up the donor site's percentile of the duration curve at the first date to be estimated. Then the same percentile is found in the target site's duration curve and the corresponding groundwater level is chosen as the estimated level at the particular date. This process is repeated for all dates available within the record of the donor sites. However, not all donor sites are given the same weight for estimation at the target site. The estimated series of groundwater levels at the target site H_t are rather put together (equation 5) by weighting each source site's contribution based on the Euclidean distance d_t to the target.

247
$$H_t = \sum_{j=1}^n w_j H_{sj} / \sum_{j=1}^n w_j$$
(5)

248 The weights are calculated based on a dissimilarity measure:

249
$$w_j = \frac{1/d_t}{\sum_{j=1}^n 1/d_t}$$
(6)

Groundwater levels are also estimated at each target site using a straightforward nearest neighbor method (NN). Here, NN means that no duration curve is reconstructed but only the actual time series of each source site L_{ti} is used, however, weighted according to eq. 5 and 6.

253 **2.5 Evalutation of Time Series Estimation**

254 The performance of the daily groundwater level prediction was evaluated using leave-one-255 out cross validation as performed by Shu and Ouarda (2012). Using a leave-one-out cross 256 validation procedure means that one (target) site is considered unmonitored and thus left out from 257 the dataset. With the remaining data set (n-1 sites), the groundwater level time series are 258 estimated at the target site. Here, a maximum of n=20 sites were allowed as donor sites. Then, the 259 performance at that site is evaluated by calculating the Kling-Gupta Efficiency (KGE), Pearson 260 correlation coefficient (R), and Root-mean-square error (RMSE) as goodness of fit measures 261 between estimated and observed time series. These steps are repeated at each of the *n* sites and the 262 average (cross-validated) estimate is found by aggregating the goodness of fit-estimates from each 263 sub-sample.

264 **2.6 Plausibility Analysis of Models**

To examine the plausibility of models used to predict percentile points along the HDC, the impact on model output is analyzed using standardized regression coefficients (MLR) and Shapley Additive Explanations values (SHAP) for XGB (Lundberg et al., 2020) using the *R* package *SHAPforxgboost* (Liu & Just, 2021). SHAP values quantify how much individual predictors, across the predictor's value range, contribute to the output variable (here the percentile point). More specifically, the SHAP value gives the difference in the model output depending on if the model is fit with or without the predictor. Using scatterplots, SHAP values can then be interpreted locally which allows understanding of the dependence structure within each model for each predictor. Further, mean absolute SHAP of all data points for each model is estimated, yielding global feature importance across each percentile. This supports understanding of the dynamic changes of importance of controls across different aquifer states and allows qualitative comparison to standardized regression coefficients of MLR models.

277 **2.7 Cluster Analysis**

278 In order to get a better understanding of the dataset, regarding similarities in dynamics and 279 subsequently site descriptors, hierarchical cluster analysis was performed. Prior to cluster analysis, 280 the selected groundwater level time series are transformed to z-scores. As input into the clustering 281 algorithm, Euclidean pairwise distances between time series were computed. Subsequently, 282 hierarchical cluster analysis using Ward linkage is performed on the matrix of pairwise distances. 283 The hierarchical relationship between the series can then be displayed in a dendrogram. From the 284 dendrograms a scree plot is constructed, by sorting the heights of the dendrograms branches and 285 plotting these against the number of nodes. The inflection point of the scree plot is then identified 286 to select the number of clusters that sufficiently describes the patterns of member time series, while 287 still generalizing the data set to a managable level.

288

289 **3 Results and Discussion**

290 **3.1 Hydrogeological Description of Clusters**

291 Cluster analysis of the data set based on similarity of groundwater level time series results 292 in hydrogeologically meaningful groups. The six identified clusters (see SI, Figure S1-S2) are 293 either made up of wells exclusively located in alluvial deposits or in alluvial deposits and outwash 294 plains. Further, cluster separation can be linked to differences in distance to stream, depth to water 295 table, size of aquifer, local hydrology and geographical location.

Figure 2A and B show that groundwater level time series in clusters C1 and C6 have similar groundwater regimes. Time series in C1 show a relatively fast response (flashy) and overprinting 298 of high peaks to varying degree, which is seen to a slightly lesser degree in C6. Inter- and intra-299 annual patterns are mostly absent. Groundwater levels in these two clusters are shallow (75% < 5)300 m) and with the wells relatively close to groundwater basin boundaries and streams in medium 301 size aquifers (Figure 2D). Presumably, these clusters represent wells tapping mainly local 302 groundwater flow systems (Giese et al., 2020). The pronounced flashiness is linked to interaction 303 with streams (E. Haaf et al., 2020) and can also be seen in the low percentiles of the duration curves 304 that are significantly steeper in the flashier C1 and C6 than other clusters (Figure 2B). Differences 305 between C1 and C6 can be attributed to the different geographical areas, with C1 located in more 306 extensive aquifers far downstream of the headwater catchment in the South and C6 located mainly 307 in smaller alluvial aquifers in the Salzach and Inn catchments at the foot of the Alps (Figure 2C 308 and SI, Figure S3).

Flashiness in cluster C2 is like C6, however, exhibiting intra-annual variations and weak inter-annual seasonality. Like C1 and C6, C2 is characterized as local flow due to the very shallow wells, however, wells are in intermediate locations in large aquifers. Therefore, dynamics are not closely coupled to the major rivers, which are at larger distances, but presumably to (unmapped) smaller creeks and to vegetation considering the shallow groundwater table.

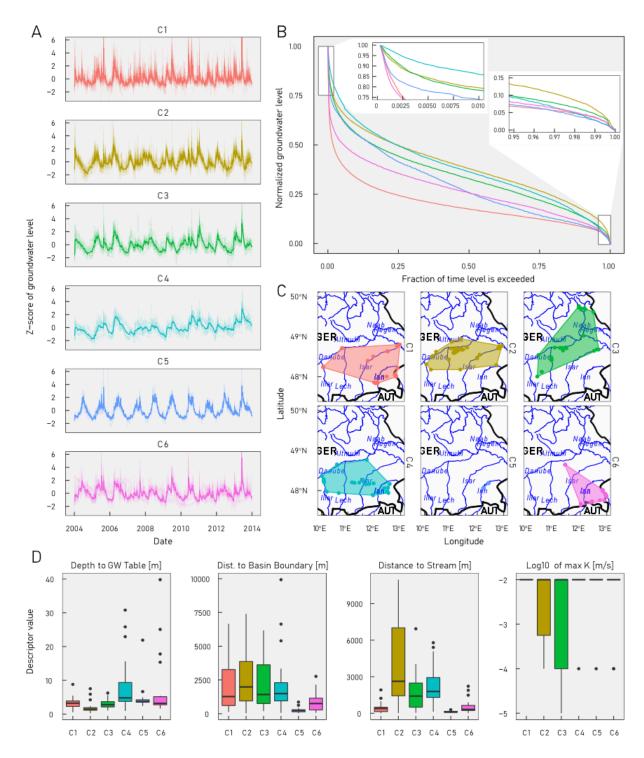
C3 is less flashy than C2, but shows a similar inter- and intra-annual pattern, which can also be seen in the similarity of the two cluster's head duration curves (Figure 2B). C3 wells are, similar to C2, located in larger aquifers, but are deeper and closer to streams, likely representing local and intermediate flow systems.

C4 has dominant inter-annual variability, which is linked to the larger distance to groundwater level and streams (E. Haaf et al., 2020). The larger inter annual variability in C4 is also seen in the less steep lower percentiles of the duration curves (Figure 2B) and is linked to mainly intermediate and regional flow systems.

322 Groundwater hydrographs in cluster C5 show a very distinct pattern compared to the 323 remaining clusters. The HDC falls steeply at lower percentiles, following the flashier C1 and C6, 324 until stabilizing and resembling more the weakly intra-annual dominated HDCs of C2 and C3, 325 before crossing back to C1 and C6 at higher percentiles, due to cluster's weak intra-annual 326 periodicity. The distinct pattern and in-group similarity of the 14 wells in C5 is explained by their

327 locations, concentrated near the Inn, which is regulated by run-of-the-river hydroelectric plants328 with pondage (Figure 2C).

329



330

331 Figure 2. A. Time series within each cluster. B. Mean of groundwater level duration curve of

332 color related to cluster in A. C. Location of cluster members with convex hull and stream

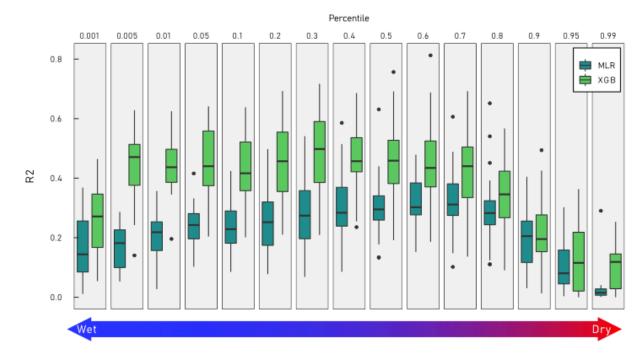
network, ISO 3166-1 alpha-3 country codes. D. Hydrogeological descriptors of sites within
each cluster.

335 3.2 Performance of HDC reconstruction

336 After regression analysis, models were found for all fifteen fixed percentage points. 337 Regression models fitted on 30 different sets of hold-out data resulted in a distribution of results 338 that are robust with regard to central tendency. Median XGB model performance on hold-out data expressed as R² is around 0.5, except for the lowest and upper percentiles (0.1%, 80-99%), i.e., 339 340 wet and dry states, where goodness-of-fit declines (Figure 3). A lower fit at the extremes is 341 expected since fewer data points make these values less robust compared to other percentiles. XGB 342 models perform significantly better than MLR models that show a similar behavior across percentiles but with lower goodness-of-fit (median R²: 0.3). Figure 3 also shows that the range of 343 R^2 is large, which is very likely related to the size of the data set. The consequence of small data 344 345 sets, when using hold-out data is that the evaluation data (here, n=32) may not be representative 346 of the training data across sets of hold-out data. Further, when running models on the entire data 347 set (training+evaluation), both XGB and MLR models show around 100% and 70% performance improvement from median R². Performance loss across hold-out data and against the entire data 348 349 set indicates that generalization from the training set is moderate and likely to improve with larger 350 data sets.

351 When comparing results to studies using an analogous methodology in streamflow, model results of R² between 0.72 and 0.99 are reported and analogous lower values in the extremes 352 353 (Mohamoud, 2010; Shu & Ouarda, 2012). This study's performance is nearly 100% higher, 354 however, neither hold-out data, cross-validation methods, or p-value adjustment for stepwise MLR 355 is used. This means that models presented in these studies are likely overfitting and generalization 356 outside of the data set could be questioned. The performance achieved on evaluation+training data 357 by XGB and MLR models in this study would thus be more comparable and are in fact in parity 358 with performance reported in streamflow studies.

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361 **3.3 Dynamic Controls on Groundwater Levels**

359

362 Relative predictor importance across percentage point models stratified by predictor class 363 for MLR and XGB models respectively is shown in Figure 4. Standardized regression coefficients 364 in MLR give both relative predictor importance (higher absolute value) but also the direction of 365 the relationship between predictor and output variable (percentile level of HDC) through the sign 366 of the coefficient (Figure 4A). Mean absolute SHAP value on the other hand, shows only relative 367 predictor importance (Figure 4B). Further, for clarity of presentation, only the most salient 368 variables are shown (MLR: variables are shown that are selected in at least 30% of hold-out data 369 sets; XGB: only the top two predictors are shown per predictor class (based on overall mean 370 absolute SHAP value).

The main result is that the importance of predictors varies across percentiles. This implies that different site (or system) descriptors to varying extents control the groundwater dynamics when the aquifer is moving from "wet" to "dry" states and vice versa. An example is distance to stream that is important through all aquifer states but dominating in wet states (both MLR and XGB, Figure 4A-B). Depth to the groundwater table, on the other hand, becomes more dominant when the aquifer is in dry states (only XGB, Figure 4B). A pattern that can be seen across all variables is that predictor strength declines significantly (approaches zero) at higher percentiles,
which is also connected to lower goodness-of-fit at these percentiles (Figure 3). Consequently,
predictability of percentiles coupled to groundwater drought is low.

380 Another important finding is that many of the most important predictors are consistently 381 selected across both MLR and XGB as well as show a similar importance progression across 382 percentiles (distance to stream, well elevation, average annual precipitation, broadleaved Forest 383 and regional slope skewness). This means that many of the important variables have a sufficiently 384 linear relationship with percentiles of groundwater head duration curves so that it can be picked 385 up by MLR. For instance, MLR models show that percentage points of the HDC increases with 386 distance to stream (the further away from streams, the less flashy the groundwater level). This is 387 plausible and expected, since streams are the aquifer's given drainage boundary and known 388 through previous regional scale empirical studies (e.g. Boutt, 2017; Giese et al., 2020; E. Haaf et 389 al., 2020; Vidon, 2012). However, SHAP values of individual data points related to XGB 390 prediction allows us to look more closely at linearity of relationships between HDC and predictor 391 value ranges (Figure 5). The SHAP values reveal a more complex relationship, where the 392 relationship between distance to stream and dynamics is constant up to about 500 m distance, 393 turning into a linear relationship, where groundwater dynamics become less flashy with distance until reaching a plateau at about 3000 m distance. Here, presumably a decoupling between 394 395 groundwater and stream occurs and a constant contribution to the HDC is reached (Figure 5). This 396 effect is consistent across aquifer states, however weakens, when the groundwater level drops into 397 dry states. The nonlinearity of relationships with threshold effects is common, as described below 398 for variables selected in Figure 5:

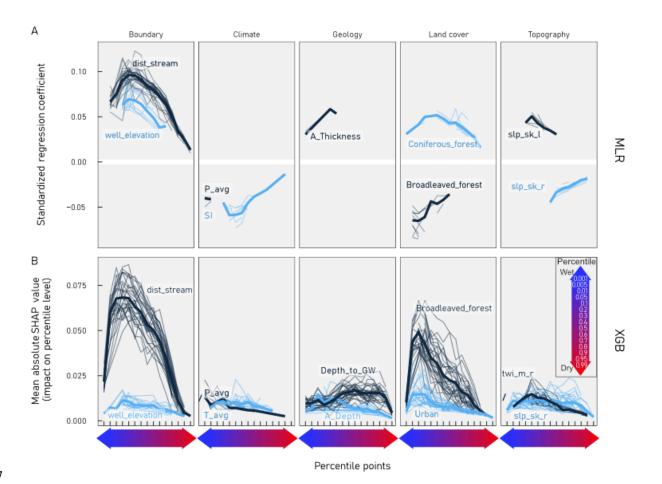
Average annual precipitation has relatively low impact on the HDC, which is also true for
 other climate predictors in this study. However, precipitation below approximately 800 mm
 leads to slightly less flashy dynamics in wet states. This can be coupled to less infiltration
 and recharge events. At higher precipitation rates, no systematic impact on HDC can be
 seen.

Depth to groundwater table only affects the HDC when very shallow, approximately 2 m
 and above. Shallow water tables increase the percentile level accordingly, meaning that
 less flashiness may be expected. Sites, where groundwater levels are very shallow may be
 coupled to discharge zones. Here the aquifer is continuously replenished through recharge

from uplands with significant upward hydraulic gradients (Gribovszki et al., 2010; Winter,
2001). Generally, this effect increases in importance at higher percentiles, i.e., in a drier
aquifer state

- If the percentage of broadleaved forests exceeds approximately 10%, groundwater levels
 become flashier in wet states, which can be linked to higher soil moisture, preferential flow
 and recharge than other land cover types, reducing surface runoff (Brinkmann et al., 2019;
 Dubois et al., 2021).
- If regional slopes are right skewed, sites are located in alluvial valley bottoms at the fringes
 of higher hill ranges (Ezra Haaf et al., 2020; Montgomery, 2001). In these locations
 amplitudes are expected to be higher due to front slope flow and mountain block recharge,
 which is also seen here particularly in wet aquifer states with lower SHAP values at higher
 slope skewness. Low slope skewness (<.0.3) on the other hand contributes to less flashy
 groundwater dynamics.

421 Overall, the progression of controls have implications not only for prediction but also 422 conceptual understanding of groundwater dynamics in this region. The nonlinear relationships of 423 groundwater dynamics and controls and the alternating dominance of these controls throughout 424 different aquifer states are likely of interest, when studying e.g., vulnerability to drought events 425 and climate change. Certainly, there is a need for a dedicated analysis of the dependence of controls 426 on aquifer states, which was outside of the scope in this study.



427

Figure 4. Relative predictor importance across percentage point models stratified by predictor class for MLR and XGB models (scales not comparable). Data from all hold-out datasets are plotted and fitted with a local polynomial regression to emphasize the central behavior of the data. A. Standardized regression coefficients show both relative predictor importance and direction of relationship between predictor and model output. B. Mean absolute SHAP value shows relative importance through impact on the output variable.

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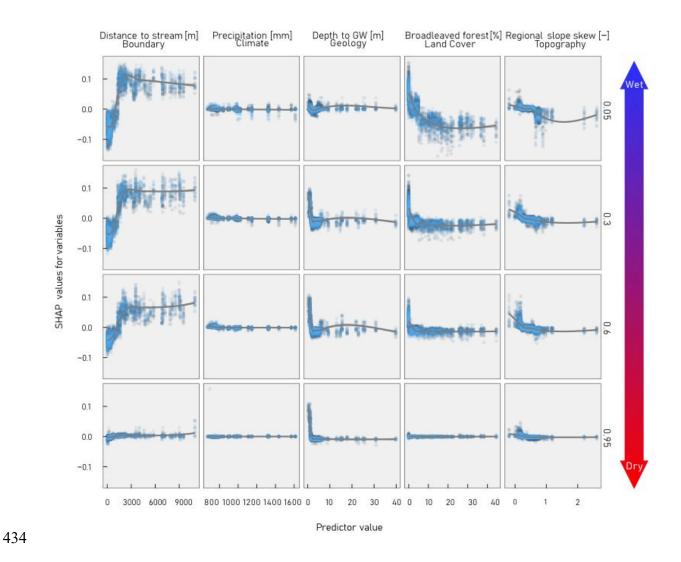


Figure 5. Relationship between feature value and impact on prediction for five selected
variables across four percentiles. Each point represents an observation of the predictor
variable and its SHAP value. Data from all hold-out datasets are plotted and fitted with a
local polynomial regression to emphasize the central behavior of the data.

439

440 **3.4 Performance of estimation techniques**

Daily groundwater level time series were estimated at target sites, using representative
models from each of MLR and XGB models as well as using the Nearest Neighbor method (NN).
The XGB model had a higher KGE than NN at 120 of 157 (76%) sites, and a higher KGE than
MLR at 136 of 157 (87%) sites. In consequence KGE is also significantly higher for XGB than
NN and MLR (Figure 6A). Interestingly, MLR has a lower median KGE than NN, (slightly higher

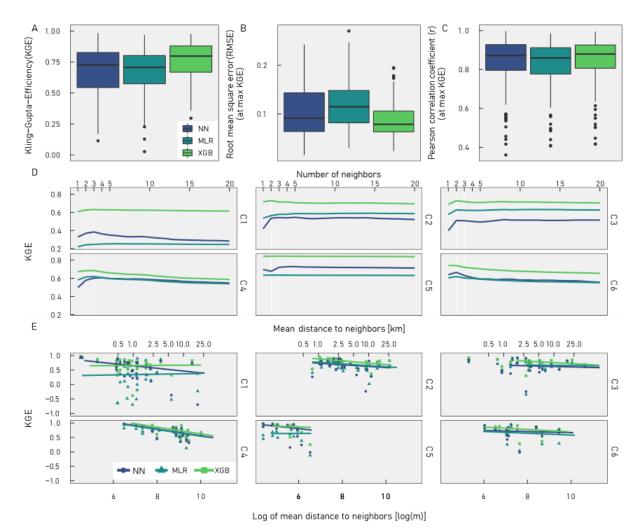
446 performance at the lower quartiles) which means that HDC modelling in the case of MLR447 deteriorates estimation on average, compared to the simple NN approach.

448 The higher performance of XGB can almost entirely be attributed to smaller amplitude 449 errors between simulated and observed time series. Amplitude errors are expressed by the RMSE 450 component of KGE, which is significantly improved when using XGB compared to NN and MLR 451 (Figure 6B). The correlation component of the KGE on the other hand shows no significant 452 differences between methods, meaning that timing errors between observed and simulated time 453 series are not significantly improved through XGB or MLR (Figure 6C). As discussed by 454 Mohamoud (2010), timing errors are coupled to the mismatch of time sequence in hydrograph 455 events (here, e.g., recharge events) at donor and target sites. Still, from a water resources 456 management perspective, the HDC estimation approach using XGB implies better estimation of 457 the quantitative status of groundwater resources through significantly reduced amplitude errors.

458 Figure 6D shows that an optimal number of donor sites (neighbors) is generally reached 459 with only 1-3 neighbors, as expressed by the maximum KGE. Sourcing more neighbors generally 460 results in plateauing or even decrease of estimation performance across different groundwater 461 regimes, as expressed by clusters C1 - C6. Although the number of optimal donor sites is 462 consistent, C4 and C6 exhibit a sharp decline, when more than three or two source sites 463 respectively are added. A possible reason for this is that these two clusters contain sites with 464 significantly deeper groundwater tables (Figure 2D). This means that source sites with e.g., more 465 shallow water table and therefore deviating groundwater response will be weighted in and cause a 466 mismatch of time sequence, decreasing the quality of the predicted groundwater level time series 467 at the target site.

468 Not only hydrogeological suitability of donor sites is important, but also proximity (Figure 469 6E). Performance decreases approximately with the natural logarithm of mean distance of 470 neighbors. However, even at large mean distances to source sites (e.g. > 5 km), estimation 471 performance at many sites may remain high. This is particularly the case for cluster C2 and C3. 472 These cluster also show significantly higher performances by both HDC-based estimation 473 techniques MLR and XGB. On the other hand, at sites with sufficient neighbors nearby (< 5 km), 474 NN is preferred over MLR. Overall, however, XGB yields best performance independently of 475 mean distance to neighbors.

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Figure 6. **A.-C.** Performance of estimation of daily groundwater level time series for the three approaches across all unmonitored sites, measured as KGE (A), RMSE (B), Pearson's r (C). **D**. Mean performance – measured as KGE - of the three estimation methods plotted against number of included neighboring sites, stratified by cluster. **E**. Performance of all sites – measured by KGE - plotted versus mean distance to neighbors, stratified by estimation method and cluster.

481 **3.5 Hydrogeological Controls and Plausibility of Models**

482 From a hydrogeological perspective, there are obviously missing descriptors to describe 483 groundwater levels, such as aquifer properties, transmissivity and storativity. These are often not consistently available at the scale of this study (regional scale), or only with a low level of certainty 484 at the level of 1-2 orders of magnitude (e.g., hydraulic conductivity in this study). However, it can 485 486 be argued that the importance of storativity in this study is reduced, since normalization on a 0-1 487 scale of groundwater level time series reduce the importance of amplitude. Regarding hydraulic 488 conductivity a relatively homogenous selection of sites is used (Figure 2D). When assuming order 489 of magnitude similarity of hydraulic conductivity, the predictor aquifer thickness (A thickness)

490 may be considered a rough proxy. With these simplifications and proxy variables, model fits are 491 acceptable, but still contain significant uncertainty, resulting in lower quality of time series 492 prediction. Adding hydraulic properties, i.e., storativity values and less uncertainty regarding 493 hydraulic conductivity to the set of predictors would likely improve the fit of regression models. 494 It would further allow for use of more heterogeneous data sets. Different strategies to extract such 495 hydraulic properties at wells from groundwater level time series of unconfined aquifers was 496 recently proposed using transfer function noise models (Peterson & Fulton, 2019) and spectral 497 analysis (Houben et al., 2022).

498 Apart from the missing hydraulic properties, other factors likely also play a role in 499 explaining the moderate goodness-of-fit of the HDC models. Some of the uncertainty may be due 500 to different hydraulic properties stratified within the zone of fluctuation. This is the case at only a 501 few sites according to the borehole logs. Other sources of uncertainty may be found in data 502 (groundwater level measurements, spatial resolution of DEM and climate data) or method of 503 estimating physiographic and climatic descriptors.. Other reasons may be found in the 504 overrepresentation of relatively shallow allowial aquifers, particularly in the north-east of the study 505 area. Using mean squared error as a loss function, regression models tend to better represent the 506 bulk of the sites within the data set, which are mainly lowland riverine aquifers with shallow 507 groundwater levels (local groundwater flow) and less so the peri-alpine river valleys in the north-508 east. A functional stratification of the data prior to HDC model building by e.g., the dominating 509 predictor distance to stream, or more conceptually-based, using the hydrological landscape concept 510 (Winter, 2001) may improve the predictive performance of the HDC models for sites that are less 511 well represented. Using these functional pre-classifications should also improve transferability of 512 methods to other study domains. For such an exercise, however, a data set would be necessary 513 with sufficient data points that ensures robust models in each functional stratum.

514

3.6 Improvement of Donor Selection

The bias of the models towards well-represented hydrogeological settings as described above, also has consequences on donor-based reconstruction of time series at unmonitored sites. As discussed in section 3.4, differences in timing error between the three methods, NN, MLR and XGB, are very small and related to the similarity of time sequences between target and donor sites. A mismatch occurs, when inadequate donor sites are selected, which can be seen for example in 520 cluster C4 and C6 (Figure 6D). Performance in these clusters declines with each additional donor 521 and is presumably related to donors for intermediate/regional flow (C4) target sites being selected 522 from (C6) sites that are located near rivers. In other words, donor sites have hydrological responses 523 that differ from the target sites. Similar responses at sites with intermediate and regional flow 524 systems can however be expected even at larger distances (Giese et al., 2020; Haaf & Barthel, 525 2018). In consequence, careful selection of donor sites is crucial to the performance of the method 526 (also pointed out by authors applying the approach to streamflow: e.g., Hughes & Smakhtin, 1996; 527 Shu & Ouarda, 2012; Smakhtin, 1999) and geographical proximity should not always be the main 528 or sole selection criteria for source sites.

529 Likely, a cleverer approach than solely proximity for donor site selection, would surely 530 improve the performance of the presented approach significantly. Such a strategy could be based 531 on a distance metric that uses physiographic and climatic site descriptors for quantification of 532 similarity between sites, as proposed for streamflow by Shu et al, 2012. However, after studying 533 the nonlinearity of relationships between site descriptors and groundwater dynamics, a non-534 continuous approach may be more useful. Often, step changes could be seen, which indicates that 535 a discrete classification approach may provide a more optimal pool of donor sites. Such classes of 536 similar responses could be developed from the SHAP values in Figure 5, for example, that neighbors must be within the same distance to stream, i.e., within one of three classes (1-500m, 537 538 500-1500, > 1500m). For many of the sites, however, nearby sites still provide the most adequate 539 timing of events. Therefore, any of the donor selection strategies discussed above must be 540 combined with an approach that applies weights to donors within the similar class based on 541 proximity.

542 4 Conclusions

543 Using the presented method, groundwater head duration curves can be transferred based 544 on comparative regional analysis of map-derivable site descriptors from monitored to unmonitored 545 sites. Neighboring donor sites can then be used to successfully reconstruct the daily groundwater 546 level time series based on the transferred duration curve. Apart from time series estimation at 547 unmonitored sites - in essence spatio-temporal interpolation - the modelling approach also gives 548 insight into hydrological processes through identification of significant controls. Specifically, at 549 the study site, controls on groundwater dynamics were nonlinear, which favors use of Machine Learning (i.e., gradient boosted regression trees) over multiple linear regression and therefore makes possible improved conceptual hydrogeological understanding as well as higher predictive skill. The method and results were robust as tested through nested cross-validation, however, require thorough testing with larger data sets for application in other hydrogeological settings.

554 The study also showed that only 1-3 neighboring donor sites are generally necessary to 555 optimally reconstruct time series of unmonitored sites. Further, the fewer nearby donor sites are 556 available, the more benefit can be drawn from using the proposed comparative regional analysis 557 approach, compared to nearest neighbor averaging of time series. Importantly, the selection of 558 donor sites was identified as a key factor to improve predictive skill and should be expanded on 559 using a combination of geographical proximity and functional classes of groundwater sites from 560 which to draw appropriate neighbors. Finally, the study shows ways forward to investigate the 561 dynamic nature of controls on groundwater levels, which may provide valuable insight to studies 562 of recharge seasonality, droughts and floods.

563 Author Contributions

Haaf conceived the study with input from all co-authors. Haaf performed the statistical analysis and wrote the manuscript. All co-authors edited and revised the manuscript and approved the final version.

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571

572 **Open Research**

573 Groundwater time series cannot be provided publicly by the authors based on the data 574 usage agreement with the LfU, but can be downloaded from 575 https://www.gkd.bayern.de/en/groundwater/upper-layer and 576 https://www.gkd.bayern.de/en/groundwater/deeper-layer. The selected station names are provided 577 in the Supplementary Information. Processed data will be made available on zenodo after

- 578 acceptance. Code for reproduction of results can be obtained from the corresponding author. All
- 579 the analysis was performed in the statistical language R (R Development Core Team, 2022) using
- apart from the packages mentioned in the body "tidyverse", "lubridate", "rsample", "vtreat," and
- 581 "selectiveInference" The authors thank the contributors of all these packages.
- 582

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